

# 1 **Uncertainty in United States Coastal Wetland Greenhouse Gas Inventorying**

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1 **Abstract**

2 Coastal wetlands store carbon dioxide (CO<sub>2</sub>) and emit CO<sub>2</sub> and methane (CH<sub>4</sub>) making them an  
3 important part of greenhouse gas (GHG) inventorying. In the contiguous United States  
4 (CONUS), a coastal wetland inventory was recently calculated by combining maps of wetland  
5 type and change with soil, biomass, and CH<sub>4</sub> flux data from a literature review. We assess  
6 uncertainty in this developing carbon monitoring system to quantify confidence in the inventory  
7 process itself and to prioritize future research. We provide a value-added analysis by defining  
8 types and scales of uncertainty for assumptions, burial and emissions datasets, and wetland  
9 maps, simulating 10,000 iterations of a simplified version of the inventory, and performing a  
10 sensitivity analysis. Coastal wetlands were likely a source of net-CO<sub>2</sub>-equivalent (CO<sub>2</sub>e)  
11 emissions from 2006 to 2011. Although stable estuarine wetlands were likely a CO<sub>2</sub>e sink, this  
12 effect was counteracted by catastrophic soil losses in the Gulf Coast, and CH<sub>4</sub> emissions from  
13 tidal freshwater wetlands. The direction and magnitude of total CONUS CO<sub>2</sub>e flux were most  
14 sensitive to uncertainty in emissions and burial data, and assumptions about how to calculate  
15 the inventory. Critical data uncertainties included CH<sub>4</sub> emissions for stable freshwater wetlands  
16 and carbon burial rates for all coastal wetlands. Critical assumptions included the average depth  
17 of soil affected by erosion events, the method used to convert CH<sub>4</sub> fluxes to CO<sub>2</sub>e, and the  
18 fraction of carbon lost to the atmosphere following an erosion event. The inventory was  
19 relatively insensitive to mapping uncertainties. Future versions could be improved by collecting  
20 additional data, especially the depth affected by loss events, and by better mapping salinity and  
21 inundation gradients relevant to key GHG fluxes.

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24 **Social Media Abstract:** U.S. coastal wetlands were a recent and uncertain source of  
25 greenhouse gasses because of CH<sub>4</sub> and erosion  
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1 **1. Introduction**

2 Managing land to optimize carbon storage and mitigate degradation is one among many  
3 strategies under consideration to curb anthropogenic greenhouse gas emissions (Griscom *et al*  
4 2017). Coastal wetlands -- defined here as salt marshes, mangroves, tidal freshwater wetlands,  
5 and tidal freshwater forests -- have received some of this attention because they can act as a  
6 net-greenhouse gas sink (Howard *et al* 2017), and because restoration (Kroeger *et al* 2017) and  
7 conservation (DeLaune and White 2012) may reduce or mitigate emissions. Regulation and  
8 market mechanisms can incentivize wetland restoration to promote emission reduction  
9 (Pendleton *et al* 2012, Wylie *et al* 2016) and myriad co-benefits (Barbier *et al* 2011, Doughty *et*  
10 *al* 2017, Griscom *et al* 2017).

11 Coastal wetlands can bury carbon (Chmura *et al* 2003, Ouyang and Lee 2013, Howard  
12 *et al* 2017) and form new soil (Morris *et al* 2002) by adding organic carbon to the soil column  
13 through sub-surface root addition (Nyman *et al* 2006). Carbon burial is a dynamic response to  
14 sea-level rise (Kirwan and Megonigal 2013, Kirwan *et al* 2016). Carbon removed from the  
15 atmosphere and incorporated into soils and plant matter is referred to throughout this paper as a  
16 'removal'. However, wetlands can also be the sources of emissions when they are eroded  
17 (DeLaune and White 2012), developed (Stein *et al* 2014), or drained for agriculture (Drexler *et al*  
18 2009). Freshwater and brackish tidal wetlands emit methane (CH<sub>4</sub>) (Bridgham *et al* 2006,  
19 Poffenbarger *et al* 2011), a more potent greenhouse gas than carbon dioxide (CO<sub>2</sub>) over the  
20 course of its atmospheric lifetime (Frolking and Roulet 2006, Neubauer and Megonigal 2015). At  
21 a national scale, in order to estimate total greenhouse gas emissions or removals, researchers  
22 need to know the areal coverage of different wetland types, the areal coverage of wetland  
23 change events, and to assign annualized CO<sub>2</sub> equivalent (CO<sub>2</sub>e) stock changes to those  
24 wetland classes and change events.

25 Spatial data, literature review, and expert assumptions are all used to inventory  
26 greenhouse gas fluxes at national scales. These inputs introduce uncertainty (IPCC 2014),  
27 which needs to be quantified to establish both levels of confidence and priorities for future  
28 research. The Intergovernmental Panel on Climate Change (IPCC) quantifies emissions and  
29 removals with 'emissions factors' and 'activities data'. For agricultural, forested and other lands,  
30 emissions factors are values assigning greenhouse gas fluxes to land cover types and change  
31 events (Eq. 1). Activities data are typically interpreted as the areal coverage of land cover type  
32 and/or land cover change events. The IPCC published guidance for national-scale greenhouse  
33 gas inventories for coastal wetlands (IPCC 2014), and the United States incorporated these for  
34 the first time in its 2017 national greenhouse gas inventory (NGGI) conducted by the  
35 Environmental Protection Agency (EPA) (EPA 2017). Our analysis is not an official part of that  
36 NGGI. Instead, we used the accounting concepts outlined therein, as well as updated literature  
37 review and spatial data, in order to improve uncertainty estimates at the national scale and  
38 highlight areas of research that could further reduce that uncertainty.

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$$Emissions\ or\ Removals\ (flux) = Activities\ (area) \times Emissions.Factor\ (flux / area)$$
  
41 1.

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43 In the NGGI, uncertainties in emissions and removals were estimated using a basic  
44 algebraic approach (IPCC 2014, EPA 2017). We address five assumptions and approaches

1 from the previous NGGI to improve uncertainty estimates in coastal wetlands: 1. The probability  
2 distributions of the activities and emissions data were not explicitly defined; 2. Key variables  
3 such as the uncertainty inherent in tidal-elevation maps were not included; 3. Uncertainties in  
4 many activities data and emissions factors are best described by non-normal distributions,  
5 which could not be accommodated using the basic algebraic approach; 4. Key assumptions,  
6 such as the depth affected by degradation events, were based on expert assessment and  
7 therefore treated as fixed values, not as probability distributions; and 5. Some inventory  
8 decisions, such as how to calculate the global warming potential (GWP) of CH<sub>4</sub> emissions and  
9 how much area to include in the inventory, have more than one recognized technique, and  
10 uncertainty from choosing among techniques was not quantified.

11 Our analysis expands upon the scope of the NGGI uncertainty analysis and explicitly  
12 identifies and quantifies uncertainty for key activities data and emissions factors. We update key  
13 datasets with new synthesis efforts (Windham-Myers and Cai in Revision) (Supplemental  
14 Information) and the results of NASA Carbon Monitoring Systems projects (Olofsson *et al* 2014,  
15 Byrd *et al* 2018, Holmquist *et al* 2018). Our research questions are: 1. How much certainty is  
16 there that CONUS coastal wetlands were a net-source or sink of GHGs from 2006 to 2011? 2.  
17 Which datasets, assumptions, or mapping categories introduce the most uncertainty into the  
18 coastal wetland category of the US national GHG inventory?  
19

## 20 **2. Methods**

21 We addressed our research questions by integrating multiple spatial and non-spatial datasets,  
22 explicitly defining uncertainty in each step, estimating total propagated uncertainty using a  
23 Monte Carlo analysis (Ogle *et al* 2003, Paustian *et al* 2006), and by quantifying the sensitivity of  
24 total emissions and removals to each input.  
25

### 26 **2.1. Time Period and 2006 to 2011 Land Cover Classes Analysed**

27 As in the NGGI (EPA 2017) we quantified area using the Coastal Change Analysis Program (C-  
28 CAP; Fig. 1; Supplemental Tab. 1). C-CAP is a Landsat-based land cover mapping product with  
29 23 land cover classes, including six types of intertidal wetlands defined by two types of salinity  
30 (palustrine and estuarine) and three types of vegetation (emergent, scrub/shrub, and forested)  
31 (NOAA 2014). We did not include seagrasses in this analysis because C-CAP's 'estuarine  
32 aquatic bed' category typically represents nearshore vegetated environments, such as kelp  
33 beds, which are not a net-carbon storing system (Howard *et al* 2017). The coastal wetland  
34 section of the NGGI inventory also did not include palustrine forested wetlands, since they fall  
35 under the purview of forested lands. We include them because information on their contribution  
36 to uncertainty is informative regardless of their reporting subcategory.

37 The NGGI inventory is required to report from 1990 to 2015, so they linearly interpolate  
38 C-CAP changes back to 1990 and forward to 2015 (EPA 2017). Although C-CAP produces land  
39 cover change maps for five-year intervals for all U.S. coastal states from 1996 to 2011, for our  
40 analysis we focus on the C-CAP 2006 to 2011 time step because it is currently the only version  
41 with accuracy assessment data. From 2006 to 2011 we mapped 240 different land cover types  
42 including, six classes of wetlands that had the same classification in 2006 and 2011, and 234  
43 types of change to, from, and between wetland classes.  
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## 2.2. Overview of Inventory Calculations

We quantified total U.S. GHG emissions and removals from coastal wetlands by mapping the area of different classes of stable wetlands and different types of change events, then multiplying that area by the summed soil, biomass and methane flux from 2006 to 2011 (Eq. 2).

$$total.flux = \sum_{i=1}^n estimated.area_i (soil.flux_i + biomass.flux_i + methane.flux_i) \quad 2.$$

In which:

$i$  is a 2006 to 2011 land cover class in  $n$  land cover classes

$estimated\ area_i$  is the total area of land cover class  $i$

Each flux is the mass CO<sub>2</sub>e emitted or stored per unit area for land cover class  $i$

As in the U.S. coastal NGGI, we defined the area of interest as the CONUS and included all C-CAP estuarine wetlands (Fig. 2) and palustrine wetlands occurring at an elevation below the highest tides. This is referred to throughout as the coastal lands definition. Since estuarine wetlands as C-CAP defines them are driven by oceanic tidal influence, we used mapped area as represented in C-CAP as fixed values (Fig. 2). Since palustrine wetlands can either be tidal or non-tidal, we used a probabilistic map of areas falling below Mean Higher High Water Spring (MHHWS) tides to map palustrine wetland area falling within the coastal zone. Palustrine wetland mapped areas were not treated as fixed values; we estimated them as a probability distribution using a mean ( $\mu_{pal,i}$ ) and standard deviation ( $\sigma_{pal,i}$ ) for each class ( $i$ ), derived from the probabilistic MHHWS map (Eq .3).

$$mapped.area_{pal,i} \sim normal(\mu_{pal,i}, \sigma_{pal,i}) \quad 3.$$

Our analysis made a distinction between mapped area and estimated area. Estimated area can be greater than or less than mapped area because unequal omission errors (errors of exclusion) and commission error (errors of inclusion) can cause a land cover class to be over- or under-mapped. We scaled mapped area by taking into account potential errors in 2011 classification (Olofsson *et al* 2014) as well as 2006 to 2011 change detection (Fig. 2.). In a simplified version of this concept, accuracy assessment matrices containing counts of true classifications and misclassifications, can be simplified down to a single estimated-to-mapped area ratio ( $r$ ) for a classification ( $i$ ) (Eq. 4). This value will be less than 1 if a land cover class is over-mapped, and greater than 1 if a land cover class is under-mapped.

$$estimated.area_i = r_i \times mapped.area_i \quad 4.$$

We estimated total emissions or removals by multiplying estimated area by the summed per area flux of soil and biomass CO<sub>2</sub> and CH<sub>4</sub> CO<sub>2</sub>e (Fig. 2). For emissions factors we treated flux data as it was reported (either positive or negative), but transformed them when necessary, so that any emissions were always represented as a negative value and removals were always represented as a positive value. For soils, if the land cover type did not change or changed but did not result in soil loss (Supplemental Information 2.3.1), then soil carbon flux was estimated as the annual soil carbon burial rate multiplied by the number of years that wetlands were present (Eq. 5). If the 2006 to 2011 class changed and represented a soil loss event, such as

1 conversion to developed, agricultural land, or open water, then emissions were estimated to be  
 2 the product of mean soil carbon density, depth lost, and fraction of that returns to the  
 3 atmosphere (Eq. 6). We quantified biomass using three vegetation classes: forested,  
 4 scrub/shrub, and emergent vegetation. We estimated biomass flux if there was a transition  
 5 between vegetation types or from vegetated to unvegetated surfaces between 2006 and 2011  
 6 (Eq. 7). We quantified CH<sub>4</sub> fluxes using two salinity classes, since freshwater wetlands  
 7 (palustrine) emit more methane than brackish to saline wetlands (estuarine) (Poffenbarger *et al*  
 8 2011). We calculated methane flux for a class by determining CH<sub>4</sub> emissions associated with  
 9 the salinity type in 2006 and 2011, summing them, and multiplying by 2.5 to normalize the flux  
 10 over 5 years (Eq. 8).

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$$12 \quad \text{soil.flux}_{no.loss} = \text{soil.burial} \times n.\text{years} \quad 5.$$

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$$14 \quad \text{soil.flux}_{loss} = -(\text{soil.carbon} \times \text{depth.lost} \times \text{fraction.returned}) \quad 6.$$

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$$16 \quad \text{biomass.flux} = \text{biomass}_{2011} - \text{biomass}_{2006} \quad 7.$$

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$$18 \quad \text{methane.flux} = -2.5 (\text{methane}_{2011} + \text{methane}_{2006}) \quad 8.$$

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## 20 **2.2. Estimating Area of Wetland Class and Change Events**

### 21 **2.2.3. Using Tide and Elevation Data to Map Coastal Palustrine Wetlands**

22 As in the previous NGGI, we mapped a subset of palustrine wetlands categorized as coastal  
 23 lands because their tidal elevation was lower than MHHWS. However, uncertainties in digital  
 24 elevation model (DEM) elevations and in mapping tidal height were not previously included in  
 25 the NGGI uncertainty analysis (EPA 2017). We enhanced the inventory by creating a  
 26 probabilistic coastal lands map (Supplemental Information: Section 2.1.).

27 For wetland surface elevation data we used DEMs that were created using Light  
 28 Detection and Ranging (LiDAR) and were aggregated by the National Oceanic and Atmospheric  
 29 Association (NOAA) for their Sea-Level Rise Viewer (NOAA 2016) (Supplemental Tab. 1).  
 30 DEMs were created to Federal Emergency Management Administration accuracy standards  
 31 (Flood 2004, Coveney 2013). DEMs have a nominal Root Mean Square Error (RMSE) of 0.185  
 32 m for low-relief areas and assume no bias (NOAA 2017). However, wetland vegetation and soil  
 33 introduce system-specific bias and random error (Chassereau *et al* 2011) not captured by the  
 34 nominal accuracy reporting. We corrected for a mean error of 0.173 m and estimate a RMSE of  
 35 0.205 m for wetland surfaces based on a weighted average of results from multiple U.S.-based  
 36 studies (Supplemental Tab. 2). We created a map of MHHWS heights using empirical Bayesian  
 37 kriging to interpolate between NOAA tide gauges. We also created a corresponding uncertainty  
 38 map incorporating random error in LiDAR mapping, datum transformations (Schmid *et al* 2013,  
 39 Leon *et al* 2014), and distance between tide gauges. We combined the DEMs, the MHHWS  
 40 map, and the associated uncertainty surfaces into a single spatial layer representing the  
 41 probability of elevation being below MHHWS (Fig. 1-2).

42 For palustrine wetlands, we treated mapped area as a random variable. For each of 111  
 43 palustrine wetland categories we extracted pixel counts by probability class for the coastal lands  
 44 map intersecting the C-CAP class and represented mapped area as a normal distribution

1 approximated from the multiple binomial distributions (Supplemental Information: Section 2.1).  
2 The means and standard deviations for all 111 palustrine wetland stable classes and palustrine  
3 wetland change events are reported in Supplemental Table 2.

#### 4 5 **2.2.4. Representing Uncertainty in Land Cover Classification and Change Detection**

6 We calculated an estimated area from mapped area (Olofsson *et al* 2014, Byrd *et al* 2018) by  
7 combining accuracy assessment matrices (McCombs *et al* 2016) with area data from C-CAP  
8 (NOAA 2014) (Supplemental Table 4-5). C-CAP did not assess classification accuracy for all  
9 individual land cover change events between 2006 and 2011. Instead there is an overall  
10 accuracy assessment for 2011 classification and one for the 2006 to 2011 generalized 'change'  
11 or 'no change' categories.

12 The accuracy assessment matrix records counts for all instances of mapped classes --  
13 what a datapoint was mapped as -- and reference classes -- what it actually was (Supplemental  
14 Tab. 4-5). We converted the accuracy assessment matrix from counts to proportional areas  
15 (Olofsson *et al* 2014, Byrd *et al* 2018), and calculated the estimated proportional area for each  
16 class as the reference class' column sum in the proportional area matrix. We used estimated  
17 and mapped area at the full map scale to calculate an estimated to mapped area ratio ( $r$ ). For  
18 each 2006 to 2011 C-CAP class, we used the appropriate  $r$  to scale mapped area by the 2011  
19 class. We then used a second  $r$  value from the 'change' and 'no change' matrix to scale again  
20 based on change detection. Additional detail on how we calculated proportional area accuracy  
21 assessment matrices and class-specific scaling factors are available in the supplemental  
22 information (Section 2.2).

23 We represented uncertainty in estimated to mapped area ratio by representing each  
24 mapped class in the accuracy assessment count matrix as a multiple multinomial distributions, a  
25 distribution that describes counts falling into two or more categories as a random variable  
26 (Supplemental Information: Section 2.2).

### 27 28 **2.3. Carbon Storage and Emissions Data**

29 As in the NGGI we calculated emissions factors for soils, and CH<sub>4</sub> based on literature review  
30 and synthesis. Unlike the NGGI we include carbon fluxes related to biomass because data is  
31 now available as part of a remote sensing calibration and validation effort (Byrd *et al* 2018), and  
32 a literature review that is part of continued inventory development (Supplemental Information:  
33 Section 2.3). We did not include N<sub>2</sub>O emissions.

#### 34 35 **2.3.1. Soil Flux Data**

36 We estimated soil carbon stock change in wetlands remaining wetlands and lands converted to  
37 wetlands as annual carbon burial rate from a literature review of lead-210 (<sup>210</sup>Pb) dated cores  
38 (Supplemental Information; Section 2.3.1). <sup>210</sup>Pb-based measurements typically integrate  
39 carbon burial over a century, compared to cesium-137 (<sup>137</sup>Cs)- and artificial plot-based  
40 measurements, which integrate carbon burial over multi-decadal to annual time scales;  
41 therefore we assumed <sup>210</sup>Pb-based rates are more representative of long-term storage rates.  
42 We described soil carbon burial using a lognormal distribution because observed removals can  
43 not be negative when strictly relying on dated sediment profiles, observed values were always  
44 greater than zero, and the data show a positively skewed distribution (Fig. 3; Tab. 1).

1 For soil carbon stock change associated with wetland loss, we used average soil carbon  
2 density values reported by Holmquist et al. (2018) to characterize the CO<sub>2</sub> emission rate (Tab.  
3 1). Holmquist et al. (2018) determined that soil carbon density did not vary significantly by  
4 depth, and that the probability distribution of soil carbon density was described well by a normal  
5 distribution, truncated so that values could not be less than zero. They also determined utilizing  
6 a single average value for all wetlands was more parsimonious and precise than stock  
7 estimates based on available maps of soil carbon.

8 The previous NGGI (EPA 2017) made two assumptions about carbon changes during  
9 wetland conversion events that were not considered in the error propagation. First, the depth of  
10 soil lost to conversion was based on a range of values reported for aquaculture and salt  
11 production pond construction (0.5-2.5m; IPCC 2014) but was fixed to 1 m. In the NGGI, this  
12 value was applied to wetland areas that converted to open water as indicated by C-CAP.  
13 Because wetland to open water conversion events were dominant in our accounting and the  
14 IPCC depth intervals for degradation were largely not applicable, we represented uncertainty  
15 regarding this assumption by using a uniform distribution ranging from 0.5 to 1.5 m (Tab. 1) to  
16 represent a wide distribution centered on 1 m. This uncertainty reflected a consensus from our  
17 coauthor group and reflected an expert assumption rather than data, as we could not readily  
18 locate or ingest any relevant data. The NGGI also assumed that 100% of the carbon released  
19 by conversion from coastal wetlands to open water is lost to the atmosphere. However  
20 (Lovelock *et al* 2017) reviewed available studies and estimated 25-50% of terrestrial carbon  
21 delivered to the marine environment was buried in ocean sediments (Baldock *et al* 2004, Cai  
22 2011, Blair and Aller 2012). Therefore we represented the fraction lost back to the atmosphere  
23 as a uniform distribution ranging from 50 to 75% (Tab. 1).

### 24 25 **2.3.2. Biomass Flux Data**

26 We utilized biomass data from (Byrd *et al* 2018) to generate emissions factors for emergent  
27 wetlands. We accounted for forested wetland biomass using a synthesis of tree diameter at  
28 breast height (DBH) for mangrove and tidal freshwater forested plots, then converting DBH to  
29 above ground biomass using allometric equations cited within the data source, or originating  
30 from a similar representative study (Supplemental Information: Section 2.3.2). We represented  
31 scrub/shrub data using a subset of the Byrd et al. (2018) biomass data, plots that were  
32 dominated by the shrub *Iva frutescens*, and a subset of the forested biomass dataset, plots in  
33 which average tree heights were lower than 5 m. We converted biomass to organic carbon  
34 using a conversion factor of 0.441 (Byrd *et al* 2018). We represented above-ground biomass  
35 with lognormal distributions because the data exhibited skewed positive distributions (Tab. 1;  
36 Supplemental Fig. 2).

### 37 38 **2.3.3. Methane Flux Data**

39 For CH<sub>4</sub> fluxes, we utilized a synthesis of annual CH<sub>4</sub> fluxes compiled by (Poffenbarger *et al*  
40 2011) and further developed as part of the 2nd State of the Carbon Cycle Report (Windham-  
41 Myers and Cai in Revision) (Supplemental Information: Section 2.3.3). Although IPCC guidance  
42 recommends separating CH<sub>4</sub> emissions by salinity class using an 18 ppt threshold (IPCC 2014),  
43 C-CAP's two salinity categories are not optimized for this purpose. We instead had to represent



1 CH<sub>4</sub> emissions with separate estuarine and palustrine emissions factors based on a 5 ppt  
2 salinity threshold (NOAA 2014) (Fig. 4).

3 We represented CH<sub>4</sub> fluxes using a normal distribution for estuarine wetlands because  
4 while the vast majority of sites indicated a net emissions scenario, one oligohaline site in New  
5 Jersey displayed net-uptake of CH<sub>4</sub> for much of the two years reported (Weston *et al* 2014) (Fig.  
6 3). We represented palustrine CH<sub>4</sub> emissions using a lognormal distribution because flux values  
7 had a skewed positive distribution and there were no instances of net-uptake of CH<sub>4</sub> (Fig. 3;  
8 Tab. 1). We estimated the global warming potential of CH<sub>4</sub> as 25 CO<sub>2</sub>e CH<sub>4</sub><sup>-1</sup> for consistency  
9 with the NGGI (IPCC 1997, EPA 2017) even though IPCC 5th Assessment Report recommends  
10 updated conversions (28 CO<sub>2</sub>e CH<sub>4</sub><sup>-1</sup> or 34 CO<sub>2</sub>e CH<sub>4</sub><sup>-1</sup> with feedbacks; Tab. 1) (Pachauri *et al*  
11 2014).

## 12 **2.4. Uncertainty and Sensitivity Analysis**

### 13 **2.4.1. Monte Carlo Analysis**

14 We propagated uncertainty using a Monte Carlo analysis (Ogle *et al* 2003, Paustian *et al* 2006,  
15 Metsaranta *et al* 2017). We calculated the inventory (Eq. 2.) 10,000 times, simulating the  
16 underlying data using random draws from the probability distributions for 145 random variables  
17 (Supplemental Information: Section 2.3.4): including normal distributions for the mapped area  
18 for each of 111 possible palustrine stable and change classes (Supplemental Tab. 2) and  
19 multinomial distributions used to randomly draw accuracy assessment matrices for twenty-three  
20 2011 C-CAP land cover classifications (Supplemental Tab. 4), and 2006 to 2011 change and no  
21 change categories (Supplemental Tab. 5).

22 We also propagated uncertainty for nine emissions factors or emission factor  
23 components (Tab. 2). For normally distributed variables we randomly drew the same number of  
24 datapoints from literature review from the probability distribution then represented the emissions  
25 factor or component as the mean of the randomly drawn data. For uniform distributions, we  
26 randomly drew a single value. For emissions factors that were lognormally distributed we  
27 randomly redrew the underlying data as in normal distributions but represented the central  
28 tendency of using the exponentiated logmean. This choice is consistent with IPCC *Wetlands*  
29 *Supplement* guidance, however arithmetic means are often used for lognormally distributed  
30 emissions factors (Levy *et al* 2017). Because the goal of this paper is to quantify the effect of  
31 assumptions on the inventory, we repeated the uncertainty analysis using the arithmetic mean  
32 of lognormally distributed values (Supplemental information: Section 3.2; Supplemental Fig. 4).

### 33 **2.4.2. Sensitivity Analysis**

34 We performed a one-at-a-time sensitivity analysis (Metsaranta *et al* 2017), meaning we  
35 categorized sensitivity of the U.S. scale emissions and removals to assumptions, datasets, and  
36 mapping accuracies by manipulating one input at a time and recording the effect. For each  
37 random variable we re-calculated the coastal wetland total GHG emissions and removals using  
38 the 0.025 quantile and 0.975 quantile values from Monte Carlo analysis, while fixing all others at  
39 their median value. We reported sensitivity of the inventory to each input as the difference in the  
40 total flux between using the input's minimal and maximal settings.

41 The sensitivity analysis also helped test the effect of some of the fundamental  
42 assumptions. For example, CH<sub>4</sub> fluxes need to be converted to CO<sub>2</sub>e, and there is controversy  
43  
44

1 about whether to use the GWP (25 CO<sub>2</sub>e CH<sub>4</sub><sup>-1</sup>) (IPCC 2014) or the Sustained Global Warming  
2 Potential (SGWP; 45 CO<sub>2</sub>e CH<sub>4</sub><sup>-1</sup>) and Sustained Global Cooling Potentials (SGCP; 203 CO<sub>2</sub>e  
3 gCH<sub>4</sub><sup>-1</sup>) which more effectively represent the system (Neubauer and Megonigal 2015). We  
4 quantified the effect of that choice by calculating the inventory using a GWP and median values  
5 for all other inputs and then recalculated changing only the GWP to SGW/CP (Neubauer and  
6 Megonigal 2015). Also, we tested the assumption of relying on the coastal lands definition for  
7 determining how much palustrine wetland area to include in the inventory compared to a tidal  
8 wetlands definition from the National Wetlands Inventory (NWI) (Hinson *et al* 2017, Najjar *et al*  
9 2018, Holmquist *et al* 2018). For this alternative analysis, we included all C-CAP palustrine  
10 wetlands intersecting an NWI-based tidal wetlands map (Holmquist *et al* 2018) and treated all  
11 palustrine mapped areas as fixed. In the sensitivity analysis we calculated the difference in total  
12 inventory between the default settings and the NWI based mapping strategy. We also we  
13 repeated the sensitivity analysis using the arithmetic mean of lognormally distributed values,  
14 and discuss the results further in the supplemental information (Section 3.2; Supplemental Fig.  
15 5).

16

### 17 **3. Results and Discussion**

#### 18 **3.1. Initial Assessment of Estimated Area**

19 The Monte Carlo analysis combining C-CAP and LiDAR DEMs define a total area of interest  
20 with a median of 3.56 million hectares (M ha; Fig. 5). Stable wetlands were the largest category  
21 (Fig. 5) with estuarine emergent wetlands dominating (1.82 M ha), followed by palustrine  
22 forested wetlands (0.68 M ha), palustrine emergent wetlands (0.54 M ha), and estuarine  
23 forested wetlands (0.19 M ha). Of the wetlands that changed to or from other categories, loss of  
24 emergent wetlands to open water was the most dominant classification. Conversion from open  
25 water to emergent wetlands was the next most important conversion but only made up for one  
26 third of the area converted from emergent wetlands to open water. The NWI-based strategy  
27 mapped fewer palustrine wetlands, especially palustrine forested wetlands, defining a total area  
28 of interest of 2.86 M ha.

29

#### 30 **3.2. Uncertainty in the CONUS 2006 to 2011 Coastal Wetland Inventory**

31 Coastal wetlands were likely to have acted as a net-source of GHG from 2006 to 2011 (Fig. 6;  
32 Tab. 1; Supplemental Tab. 7). Across the 10,000 Monte Carlo iterations median total net-  
33 emission was -10.3 million tonnes (M tonnes) of CO<sub>2</sub>e per year (yr<sup>-1</sup>) over five years with a  
34 confidence interval ranging from -1.6 to -21.3 M tonnes CO<sub>2</sub>e yr<sup>-1</sup>. Although the confidence  
35 intervals were wide they were strictly negative, which support the conclusion of net-emissions  
36 from 2006 to 2011.

37 Separating estuarine wetlands, which have lower CH<sub>4</sub> emissions, and palustrine  
38 wetlands, which have higher CH<sub>4</sub> emissions, indicates that both classes are more likely to have  
39 acted as net-emitters (Tab. 2). However, estuarine wetlands emissions were more likely  
40 occurring due to wetland conversion events (Fig. 6). While overall stable and gaining estuarine  
41 wetlands acted as a net-sink and stable and gaining palustrine wetlands a net-source according  
42 to their median values, both categories had uncertainties spanning both net-emissions and net-  
43 storage scenarios.

44

### 3.3. The Dominant Contributions to National-Scale Uncertainty

CONUS-scale total flux was most sensitive to inputs in four major classes: uncertainty in emissions and burial data, assumptions about how to calculate the inventory, C-CAP 2006 to 2011 change detection accuracy, and C-CAP 2011 classification accuracy (Fig. 7; Supplemental Tab. 7). Overall the inventory was most sensitive to uncertainty in the underlying emissions and storage data, and to assumptions made. Uncertainty arising from the probabilistic coastal lands mapping was not a dominant contributor to total uncertainty in this framework.

Uncertainty in palustrine CH<sub>4</sub> emissions, had the greatest effect on the inventory estimates for CONUS coastal wetlands, 11.6 M tonne CO<sub>2</sub>e yr<sup>-1</sup> (Fig. 7; Supplemental Tab. 7). The average depth of soils lost to erosion, extraction, or drainage, was second most impactful and had a 9.4 M tonne CO<sub>2</sub> yr<sup>-1</sup>. Estuarine CH<sub>4</sub> emissions were also important and had a 8.5 M tonne CO<sub>2</sub>e yr<sup>-1</sup> effect. Soil carbon burial rate had a 5.2 M tonne CO<sub>2</sub>e yr<sup>-1</sup> effect and assumptions made about the fraction of soil carbon lost to the atmosphere had a 3.9 M tonne CO<sub>2</sub>e yr<sup>-1</sup> effect.

The decision to use GWP over SGWP/CP had a median effect of 8.8 M tonnes of CO<sub>2</sub>e yr<sup>-1</sup>. The alternate choice moved the estuarine stable and gains sector from net-storing (+2.2 M tonnes CO<sub>2</sub>e yr<sup>-1</sup>) using GWP to net-emitting (-2.0 M tonnes CO<sub>2</sub>e yr<sup>-1</sup>) using SGW/CP (Fig. 7; Supplemental Tab. 8). Emissions from stable palustrine wetlands overtook palustrine soil and biomass losses when using SGW/CP. The SGW/CP choice increased the estimate of total CO<sub>2</sub>e emissions 89% over the traditional GWP model.

Uncertainty in mapping also contributed to uncertainty in the inventory. 2006 to 2011 change detection was the most uncertain mapping category. Notably, we drew a different conclusion regarding the 2006 to 2011 change than the official C-CAP accuracy assessment (McCombs *et al* 2016). We concluded that change was under-mapped while McCombs *et al.* concluded change was over-mapped (Supplemental Information: Section 3.1; Supplemental Fig. 3). This occurred because McCombs *et al.* raw counts for the accuracy assessment matrix and we used a proportional area matrix (Olofsson *et al* 2014).

Sensitivity of the inventory to input uncertainty dropped precipitously for the remaining inputs. These include the decision between using a coastal lands definition to identify palustrine wetlands and the stricter NWI-based definition (2.0 M tonne CO<sub>2</sub>e yr<sup>-1</sup> effect) (Fig. 7; Supplemental Tab. 7). The effect of uncertainty in fluxes associated with changes in forested and scrub/shrub biomass and carbon density for eroded soils range from 0.6 to 0.1 M tonnes CO<sub>2</sub>e yr<sup>-1</sup>. Classification accuracy introduced uncertainty for estuarine aquatic beds, open water, unconsolidated shore and palustrine aquatic beds. In our accounting, these all indicate soil loss events.

### 3.4. Implications for Future Research

Uncertainty estimates are important components of complete and transparent GHG inventories (EPA 2017). Uncertainty information is not intended to dispute the validity of the estimates, but rather to help prioritize efforts to improve accuracy and guide future decisions. We recommend improving process models for CH<sub>4</sub> emissions and soil carbon burial, increasing the number of observations for key inputs, and developing more detailed and accurate maps for categories relevant to coastal wetland carbon cycling and inventory estimates.

### 3.4.1. Improving Process Models for CH<sub>4</sub> Emissions and Soil Carbon Burial

The uncertainty and sensitivity analysis presented herein suggest that uncertainty could be reduced at the scale of the contiguous U.S. primarily by improving data availability and process-based models for CH<sub>4</sub> emissions, CH<sub>4</sub> radiative forcing, and carbon burial rates. Net-wetland CH<sub>4</sub> emission combines CH<sub>4</sub> production by methanogenic archaea under anaerobic conditions, CH<sub>4</sub> oxidation and consumption by methanotrophic bacteria mainly under aerobic conditions, and CH<sub>4</sub> transport to the atmosphere (Conrad 1989, Whalen 2005). Major controls of these processes include: water table position; soil temperature; sulfate supply and potential production of hydrogen sulfide, a methanogen toxin, for which salinity is a proxy for; vegetation, including both biomass and species composition, which may facilitate CH<sub>4</sub> transport from soil production sites to the atmosphere; and primary production of vegetation, since new photosynthate may be a substrate for methanogenesis (Wang *et al* 1996, Walter and Heimann 2000). Large discrepancies have also been noted between chamber and eddy covariance measurements of CH<sub>4</sub> fluxes (Hendriks *et al* 2010, Krauss *et al* 2016), suggesting the need for additional comparisons between these two methods.

The use of GWPs serves an important policy need because GWPs are transparent and tractable. However, GWPs are an oversimplification because modeling CO<sub>2</sub>e in power units (W m<sup>-2</sup>) that relate directly to radiative forcing is several steps removed from actual climate impacts such as changes in temperature, precipitation, and sea level. The SGW/CP model is equally transparent and tractable, but more closely represents reality by acknowledging that changes in GHG emissions persist over several years (Neubauer and Megonigal 2015). Therefore we recommend that SCW/CP's should be considered for adoption by the IPCC. When considering the consequences of GHG inventory data beyond the IPCC context, ecosystem scientists and policy analysts should discuss metrics that are independent of time frames, such as switchover time, as they are more informative of the long-term impacts (Frolking and Roulet 2006). Our uncertainty analysis is focused on variables that are inputs to GWP and SGW/CP models, but there is an ongoing need to address the uncertainty introduced by using these models to underpin climate policy.

Currently, IPCC guidance recommends applying separate carbon burial rates to different wetland types and ecoregions to increase accuracy. However, multiple studies suggest other relevant geographic and methodological factors need to be considered in the US inventory. In some locations, accelerating sea-level rise is expanding the area conducive to carbon burial, potentially increasing carbon burial rates (Kirwan and Mudd 2012, Hill and Anisfeld 2015). A sensitivity analysis of the Marsh Equilibrium Model highlighted relative sea-level rise, plant productivity and relative tidal elevation as dominant drivers of carbon sequestration in stable wetlands (Morris and Callaway in Press). Elevation/inundation gradients were correlated with sediment accretion dynamics in San Francisco Bay (Callaway *et al* 2012). Finally, there are many ways to measure carbon burial that integrate different time scales: decades- <sup>137</sup>Cs, centuries- <sup>210</sup>Pb, or millennia- <sup>14</sup>C (Turetsky *et al* 2004). We recommend that future studies rectify the complex interactions between regional variability in relative sea-level rise, plant productivity, local elevation/inundation dynamics, and the potential effects of measuring this carbon burial using differing methods.

### 3.4.2. Increasing Data Availability for Key Inputs

1 Some inputs in the inventory could be improved by targeted studies and additional data  
2 collection, including soil depth affected by conversion to open water, and percent carbon  
3 returned to the atmosphere upon loss.

4 Available data on elevation loss due to the diking of wetlands for agriculture (Drexler *et al*  
5 *al* 2009), and the mass lost to 50 cm depth following vegetation die off (Lane *et al* 2016), are not  
6 suitable proxies for the vast majority of losses occurring from 2006 to 2011, estuarine emergent  
7 to open water conversions resulting from hurricane impacts and erosion in the Gulf of Mexico  
8 (NOAA 2014). Although average carbon mass at depth in wetland soils is well constrained for  
9 coastal wetlands (Sanderman *et al* 2018, Holmquist *et al* 2018), the sensitivity of this carbon  
10 stock to different disturbances across regions, relative elevations, and time is not well known.

11 Uncertainty in assumptions about carbon loss is not unique to this study and was  
12 discussed explicitly in a recent global analysis of soil and biomass loss from mangrove  
13 conversions (Sanderman *et al* 2018), which report that the rate and forms of carbon loss may  
14 depend on soil type and depth (Donato *et al* 2011). Because assumptions about loss events  
15 vary from study to study, and because of the fact that these assumptions are dominant  
16 contributors to uncertainty (Fig. 7), future research should prioritize empirical and modeling  
17 studies that constrain depth and percent carbon loss due to wetland conversion events.

### 18 **3.4.3. Improving Mapping Capacity of Tidal Carbon Relevant Gradients**

19 The Wetlands Supplement of the IPCC report provides two CH<sub>4</sub> emissions factors for wetlands,  
20 one for fresh to brackish conditions and another for higher salinity (18 ppt threshold)  
21 (Poffenbarger *et al* 2011, Bridgham *et al* 2013). However, C-CAP salinity categories do not  
22 match these categories, instead mapping estuarine and palustrine (5 ppt threshold; Fig. 3). This  
23 inconsistency limits our ability to confidently assess the true GHG balance for saline wetlands at  
24 the national scale. We propose developing maps and data to support at least three categories of  
25 salinity — saline (>18 ppt), brackish (0.5-18 ppt), and fresh (< 0.5 ppt) — in order to reduce  
26 uncertainty in landscape scale CH<sub>4</sub> emissions from coastal wetlands (Fig. 4).

27 Existing remote sensing approaches for vegetation and inundation dynamics could  
28 improve mapping both CH<sub>4</sub> emissions and carbon burial rates. Recent strides in mapping  
29 coastal wetland vegetation biomass (Byrd *et al* 2018), vegetation species classification  
30 (Immitzer *et al* 2016) and seasonal dynamics (Mo *et al* 2015) could provide more detailed  
31 vegetation descriptions that would be a proxy for salinity zones. For inundation/elevation  
32 regimes, extensive coastal DEMs are available, but lack the accuracy to adequately map tidal  
33 flooding depth and inundation time at relevant scales and could be improved by integrating  
34 additional remote sensing and modeling (Hladik *et al* 2013, Parrish *et al* 2014, Buffington *et al*  
35 2016). Future studies should quantify the precision needed for DEMs in the tidal zone. Currently  
36 soil emissions factors are calculated using tabular data, however improvements in mapping  
37 should be leveraged to support spatially-explicit approaches in future versions of the inventory  
38 incorporating trends in productivity and seasonality (Knox *et al* 2017), variation in carbon  
39 mineralization rates (Mueller *et al* 2018), edaphic factors and geomorphology (Rovai *et al* 2018).  
40 Many improvements may be forward-looking and hindcasting may not be appropriate (Byrd *et al*  
41 2018), and spatially-explicit approaches should only be utilized only if they actually do improve  
42 precision and accuracy of inventorying compared to simpler approaches (Holmquist *et al* 2018).  
43

1 Biomass changes were not a top contributor to uncertainty, but changes in forested and  
2 scrub/shrub biomass were the ninth and fifteenth contributors to uncertainty respectively. This  
3 study quantified the effect of uncertainty by upscaling means and uncertainties from multiple  
4 field studies, however remote sensing approaches using LiDAR, RADAR, object based image  
5 detection, and optical remote sensing, can all be used to characterize biomass changes on local  
6 to regional scales (Byrd *et al* 2018). Future studies could expand the uncertainty and sensitivity  
7 analysis to capture the effect that uncertainties in genus-specific assessments of wood density  
8 (Jenkins *et al* 2003), biomass carbon content (Byrd *et al* 2018), and the contributions and decay  
9 rates of downed wood (Krauss *et al* 2018).

10 C-CAP's accuracy was not a dominant contributor to the overall uncertainty in the  
11 inventory, but we were only able to quantify this from 2006 to 2011. C-CAP is available for the  
12 entire CONUS coastal zone from 1996 to 2011, and trends were extrapolated out back to 1990  
13 and forward to 2015 for the NGGI inventory. Future studies are needed to assess accuracy for  
14 earlier time steps.

#### 15 16 **4. Conclusions**

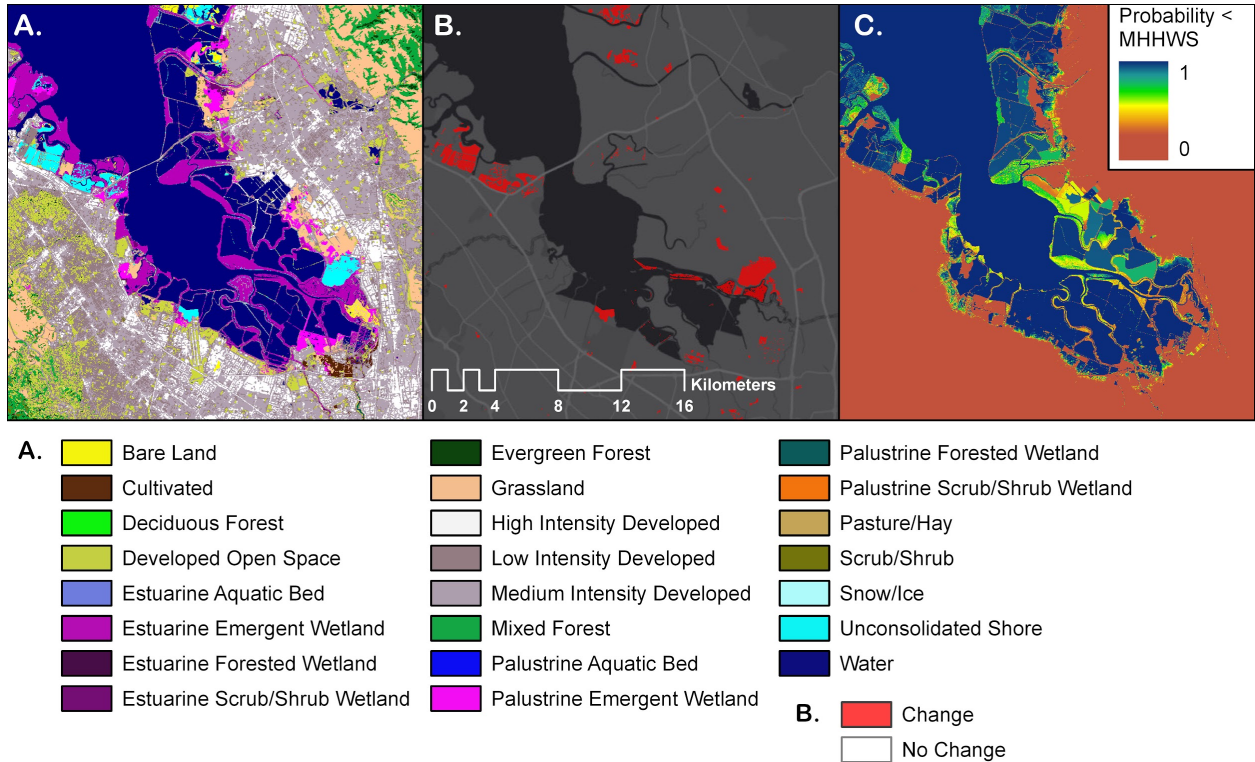
17 Uncertainty in CONUS coastal wetland greenhouse gas inventory estimates comes mostly from  
18 lack of knowledge on CH<sub>4</sub> emission variability, the fate of soil carbon post-conversion, and an  
19 inability to extrapolate trends to available map products. Switching from GWP to SGW/CP  
20 increases the overall calculation of CO<sub>2</sub>e impacts from 2006 to 2011 by 89%. The underlying  
21 mapping products, C-CAP, and the probabilistic coastal lands layer for mapping tidal freshwater  
22 wetland extent, were not dominant contributors to uncertainty. However, the inventory  
23 development could benefit from improved change detection, accuracy assessments that go  
24 back further in time, and improved mapping of intermediate salinities and inundation gradients.  
25 Our analysis provides a framework to track improvements to the coastal wetland GHG inventory  
26 as more data and improved process knowledge become available. The data used here were not  
27 collected for the purpose of the inventory; future improvements will demand targeted investment  
28 in data collection, model improvements, spatial product development, and more extensive,  
29 independent accuracy assessments.

30  
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32 primarily by NASA Carbon Monitoring Systems (NNH14AY671) and the USGS Land Carbon  
33 Program, with additional support from The Smithsonian Institution, The Coastal Carbon  
34 Research Coordination Network (DEB-1655622), and NOAA Grant: NA16NMF4630103. We  
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37 and Isa Woo, David Klinges, and Camille Stagg.

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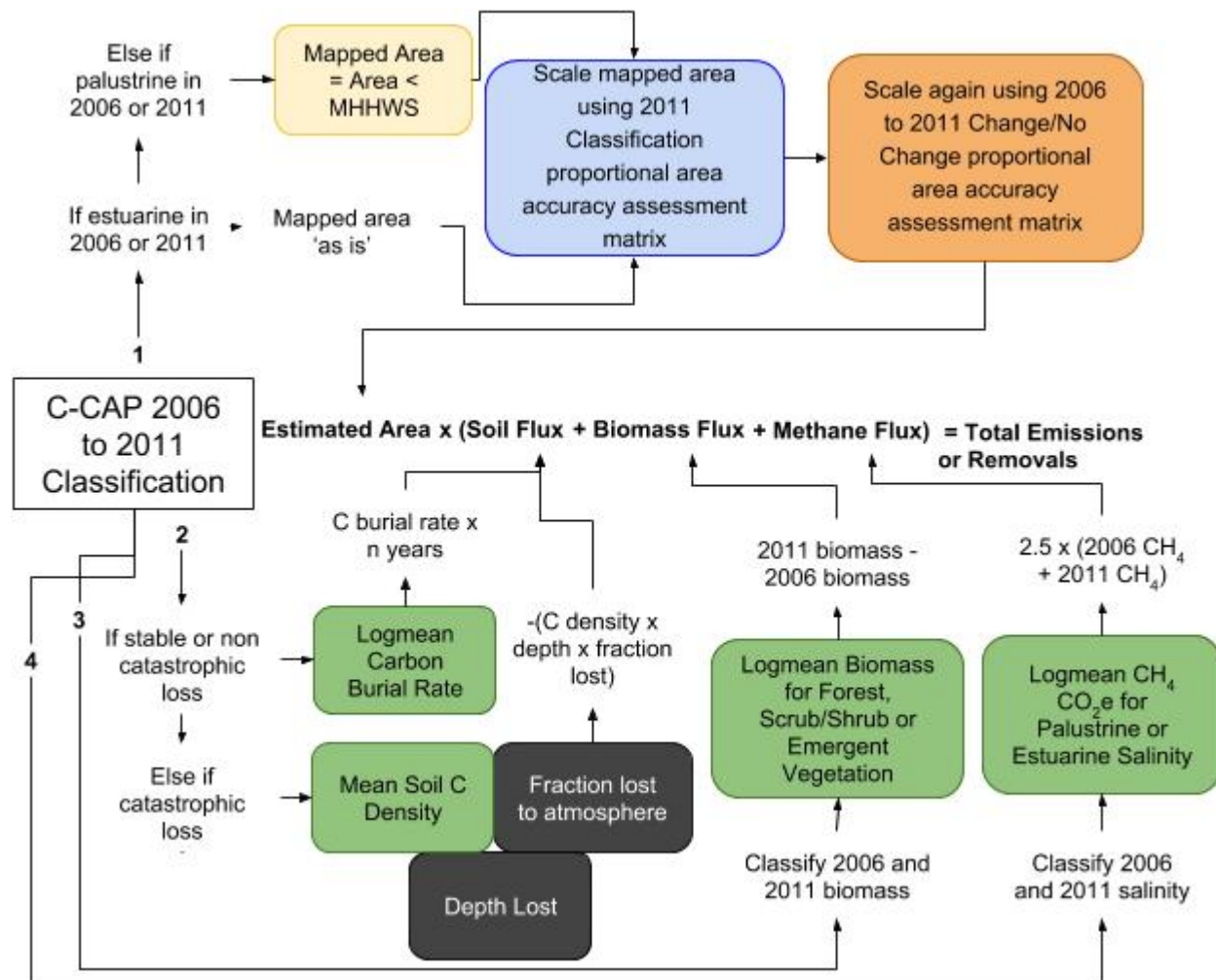
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**Figures and Tables:**



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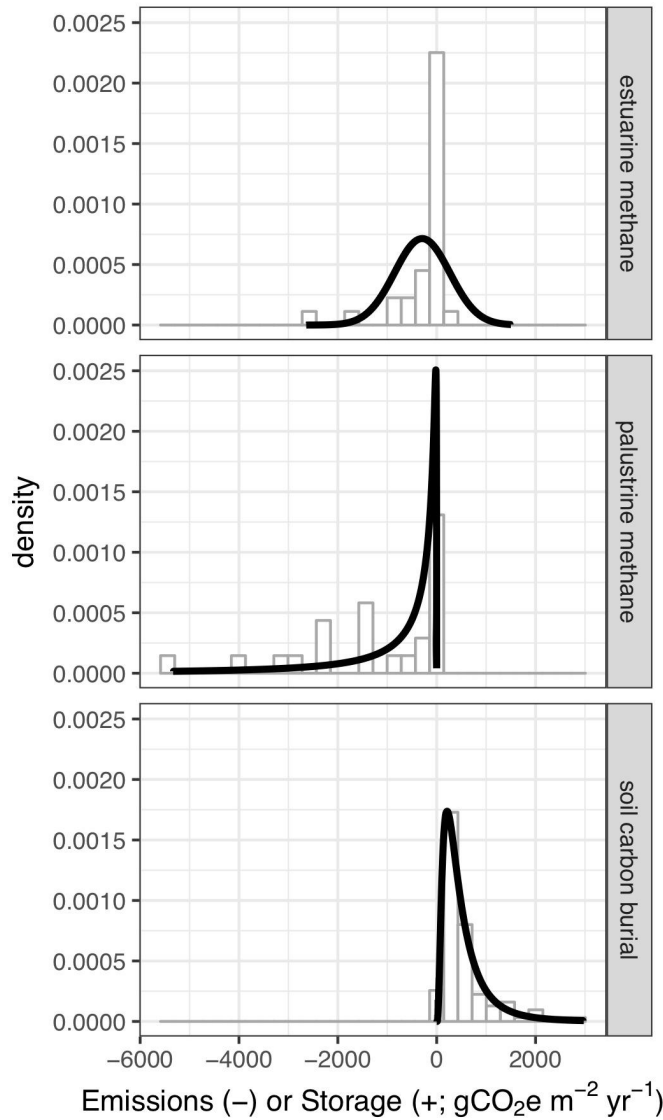
Figure 1: The three mapping layers used in our coastal wetland greenhouse gas inventory viewed for San Francisco Bay. A. 2011 Coastal Change Analysis Program (C-CAP) Land Cover Classifications. B. 2006 to 2011 C-CAP Change Map (Basemap © ESRI, permission pending). C. A probabilistic coastal lands map, showing the probability elevation is below twice highest monthly tide level, mean higher high water for spring tides (MHHWS).



1  
 2 Figure 2. Flow chart outlining how we integrated coastal wetland maps based on the Coastal  
 3 Change Analysis Program (C-CAP) land cover and land cover change products with ground  
 4 based data on soil, biomass, and methane flux. Each rounded box shows a stage at which we  
 5 quantified and propagated uncertainty. 1. How we estimated area integrating C-CAP and a  
 6 probabilistic map of area falling below mean higher high water spring tide (MHHWS) elevation.  
 7 2. How we estimated soil carbon burial and losses. 3. How we estimated biomass gains and  
 8 losses. 4. How we estimated methane emissions or removals. Colors match later categorization  
 9 of different inputs in the later sensitivity analysis (Fig. 7).

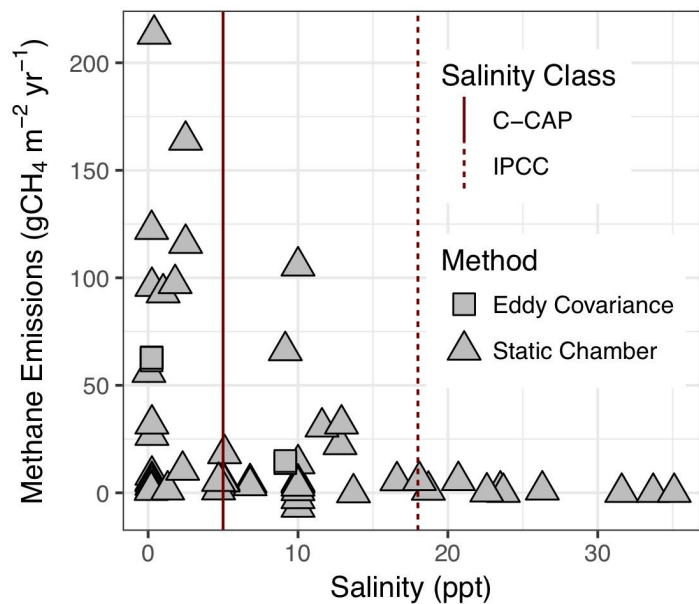
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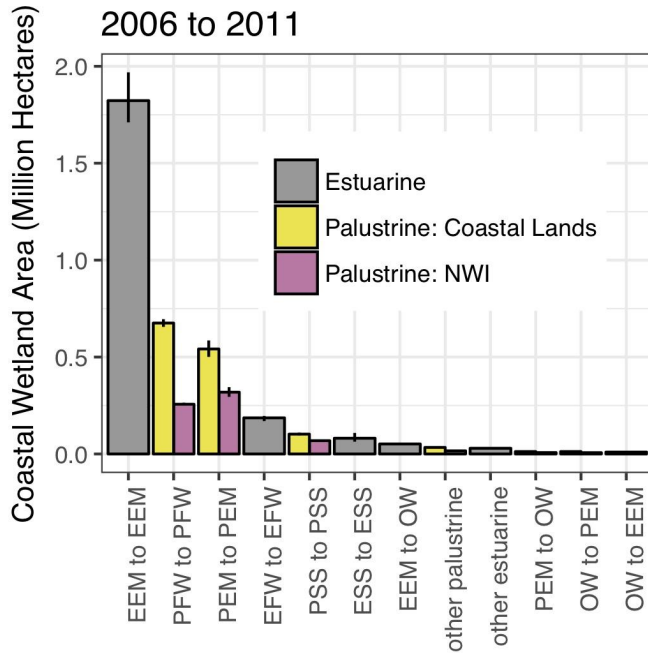


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Figure 3: Histograms and probability distributions of methane emission factors (converted to CO<sub>2</sub>e using 25x global warming potentials) and soil carbon burial rates.



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2 Figure 4: The two available salinity classes defined by the 5 ppt threshold in C-CAP are not  
3 ideal for mapping differences in methane emissions, especially when compared to the 18 ppt  
4 threshold recommended by IPCC. Data (Windham-Myers and Cai in Revision) originate from  
5 both static chamber and eddy flux covariance measurements.  
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2 Figure 5: Medians and confidence intervals for areas of wetland change classes. There are two  
 3 definitions of palustrine wetlands. the probabilistic coastal lands definition (yellow bars) and the  
 4 tidal wetlands definition based on National Wetlands Inventory (purple bars). EEM = Estuarine  
 5 Emergent Wetland, PFW = Palustrine Forested Wetland, PEM = Palustrine Emergent Wetland,  
 6 EFW = Estuarine Forested Wetland, PSS = Palustrine Scrub/Shrub, ESS = Estuarine  
 7 Scrub/Shrub, OW = Open Water.

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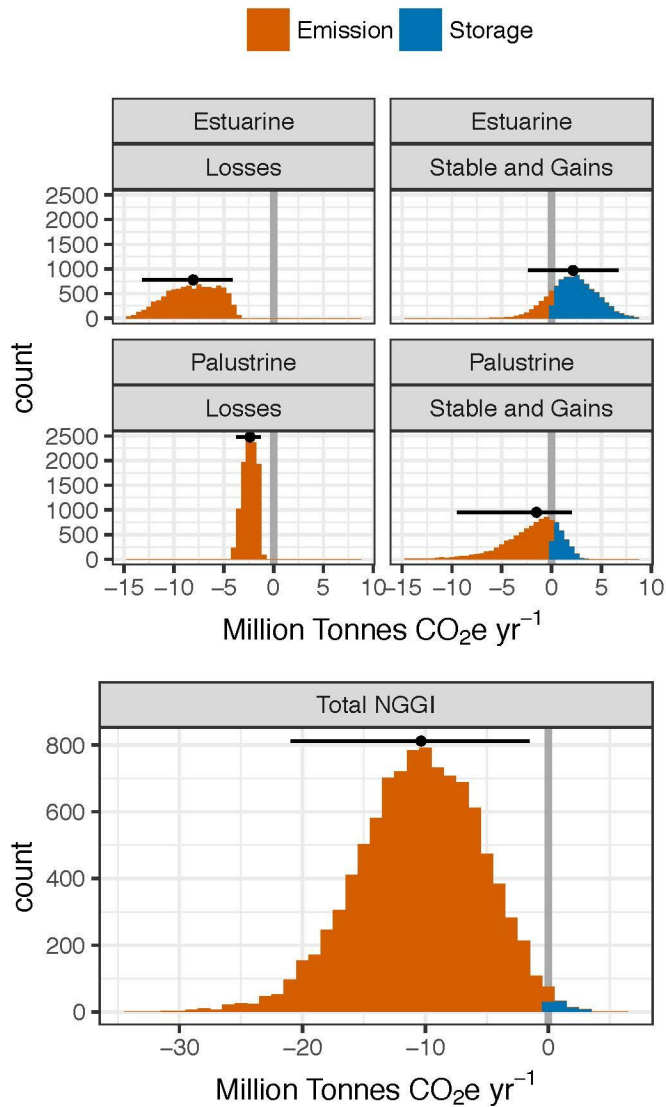
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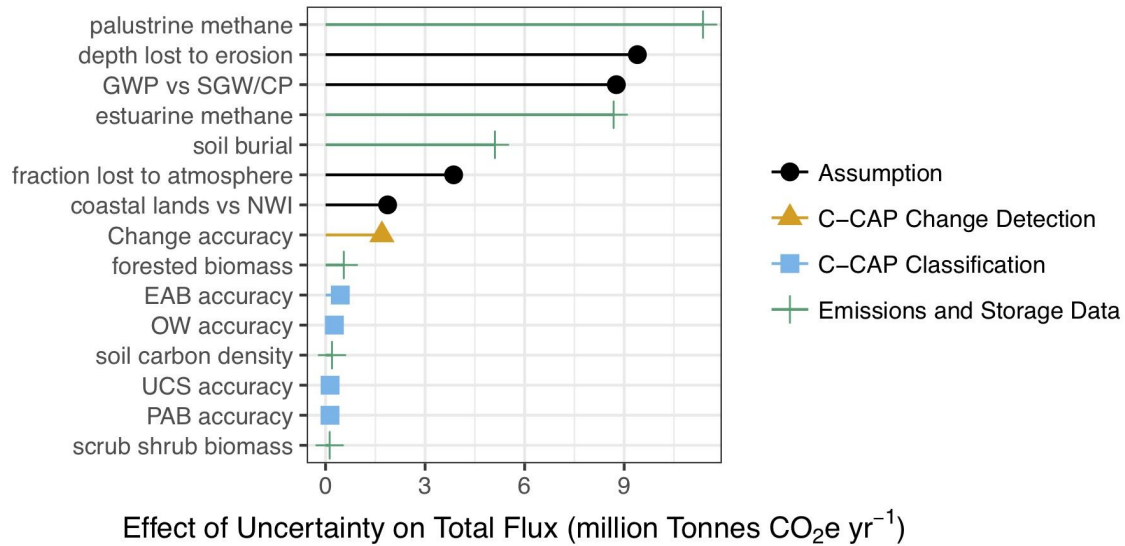
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 2 Figure 6: CONUS inventory results of 10,000 Monte Carlo simulations, shaded to distinguish  
 3 simulations resulting in a net-emission (orange) or a net-storage (blue) scenario. The thick grey  
 4 vertical line at 0 separates these scenarios. Points indicate medians, and black horizontal lines  
 5 the upper and lower 95% confidence intervals. Top panels separate fluxes from estuarine and  
 6 palustrine wetlands, and from wetlands that were lost from those that were stable or gained  
 7 area. The bottom panel shows net-annualized emissions from 2006 to 2011.

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Effect of Uncertainty on Total Flux (million Tonnes CO<sub>2</sub>e yr<sup>-1</sup>)

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Figure 7: These fifteen inputs introduced the most uncertainty into the Coastal Wetland National Greenhouse Gas Inventory (NGGI) according to a one-at-a-time sensitivity analysis. GWP: Global Warming Potential, SGWP: Sustained GWP, SGCP: Sustained Global Cooling Potential, NWI: National Wetlands Inventory, EAB: Estuarine Aquatic Bed, OW: Open Water, UCS: Unconsolidated Shore, PAB: Palustrine Aquatic Bed.

1 Table 1: Summary of probability distributions and dataset sizes used to simulate emissions  
 2 factors in the Monte Carlo analysis:  $\mu$  = mean,  $\sigma$  = standard deviation,  $\alpha$  = mean of the natural  
 3 log-transformed data,  $\beta$  = standard deviation of the natural log-transformed data, and min and  
 4 max are the minimum and maximum values of a uniform distribution.

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<b>Emissions Factor or Emission Factor Component</b>	<b>Probability Distribution</b>	<b>n</b>	<b>Moment 1</b>	<b>Moment 2</b>
Carbon Burial ( $\text{gCO}_2 \text{ m}^{-2} \text{ year}^{-1}$ )	Lognormal	109	$\alpha = 5.98$	$\beta = 1.05$
Soil Carbon Density ( $\text{gCO}_2 \text{ m}^{-3}$ )	Truncated Normal	8280	$\mu = 99000$	$\sigma = 47667$
Depth of Soil Affected by Loss Events (m)	Uniform	1	Min = 0.5	Max = 1.5
Soil Carbon Fraction Returned to Atmosphere (fraction)	Uniform	1	Min = 0.5	Max = 0.75
Emergent Biomass Change ( $\text{gCO}_2 \text{ m}^{-2}$ )	Lognormal	2345	$\alpha = 6.36$	$\beta = 1.04$
Scrub/Shrub Biomass Change ( $\text{gCO}_2 \text{ m}^{-2}$ )	Lognormal	33	$\alpha = 8.21$	$\beta = 1.97$
Forested Biomass Change ( $\text{gCO}_2 \text{ m}^{-2}$ )	Lognormal	79	$\alpha = 10.57$	$\beta = 0.75$
Estuarine $\text{CH}_4$ Emissions (GWP; $\text{gCO}_2\text{e m}^{-2} \text{ year}^{-1}$ )	Normal	31	$\mu = 292.10$	$\sigma = 558.21$
Palustrine $\text{CH}_4$ Emissions (GWP; $\text{gCO}_2\text{e m}^{-2} \text{ year}^{-1}$ )	Lognormal	24	$\alpha = 6.10$	$\beta = 1.80$
Estuarine $\text{CH}_4$ Emissions (SGWP/SGCP; $\text{gCO}_2\text{e m}^{-2} \text{ year}^{-1}$ )	Normal	31	$\mu = 477.87$	$\sigma = 1061.80$
Palustrine $\text{CH}_4$ Emissions (SGWP/SGCP; $\text{gCO}_2\text{e m}^{-2} \text{ year}^{-1}$ )	Lognormal	24	$\alpha = 6.69$	$\beta = 1.80$

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Table 2: Medians and confidence intervals for CONUS coastal wetland emissions (-) and storage (+) from 2006 to 2011 in million tonnes (Teragrams) of CO<sub>2</sub>-equivalent (CO<sub>2</sub>e) per year.

<b>Land Cover Change Type Analysed</b>	<b>lower confidence interval (0.025)</b>	<b>median (0.5)</b>	<b>upper confidence interval (0.975)</b>
Estuarine Losses	-13.3	-8.1	-4.1
Estuarine Stable and Gains	-2.3	2.2	6.7
Palustrine Losses	-3.7	-2.4	-1.3
Palustrine Stable and Gains	-9.6	-1.5	2
Total	-21.3	-10.3	-1.3

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