

1 **Interacting effects of wildlife loss and climate on ticks and tick-borne disease**

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16 Running Head: Defaunation and climate effects on ticks

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18 *Rickettsia*

19

20 **Abstract:**

21 Both large-wildlife loss and climatic changes can independently influence the prevalence and  
22 distribution of zoonotic disease. Given growing evidence that wildlife loss often has stronger  
23 community-level effects in low-productivity areas, we hypothesized that these perturbations  
24 would have interactive effects on disease risk. We experimentally tested this hypothesis by  
25 measuring tick abundance and prevalence of tick-borne pathogens (*Coxiella burnetii* and  
26 *Rickettsia* spp.) within long-term, size-selective large-herbivore exclosures replicated across a  
27 precipitation gradient in East Africa. Total wildlife exclusion increased total tick abundance by  
28 130% (mesic sites) to 225% (dry, low-productivity sites), demonstrating a significant interaction  
29 of defaunation and aridity on tick abundance. When differing degrees of exclusion were tested  
30 for a subset of months, total tick abundance increased from 170% (only mega-herbivores  
31 excluded) to 360% (all large wildlife excluded). Wildlife exclusion differentially affected  
32 abundance of the three dominant tick species, and this effect varied strongly over time, likely due  
33 to differences among species in their host associations, seasonality, and other ecological  
34 characteristics. Pathogen prevalence did not differ across wildlife-exclusion treatments, rainfall  
35 levels, or tick species, suggesting that exposure risk will respond to defaunation and climate  
36 change in proportion to total tick abundance. These findings demonstrate interacting effects of  
37 defaunation and aridity that increase disease risk, and they highlight the need to incorporate  
38 ecological context when predicting effects of wildlife loss on zoonotic disease dynamics.

39

40 **Introduction:**

41 Zoonotic diseases are a rising concern worldwide [1–3]. Yet, amid rapidly declining wildlife  
42 populations and global climate change, there is no consensus on how these perturbations will  
43 independently and interactively affect zoonotic disease risk. Anthropogenic land-use change is  
44 likely to play a substantial role in facilitating outbreaks through a variety of mechanisms [2,4],  
45 including changes to wildlife host populations and communities [3–6]. Meanwhile, climate  
46 changes can have substantial and variable effects on zoonotic diseases [7,8], even when  
47 considered in isolation of changes to host populations. Thus, the combined effects of wildlife  
48 loss and climate change are likely to be complex [7,9], but data are lacking, especially for  
49 regions where medical resources and research efforts are low and zoonotic disease risk is highest  
50 [2]. Although there has been a widespread call for more research on the net effects of  
51 anthropogenic changes on disease and disease vectors globally [3–5], large-scale experimental  
52 tests remain scarce.

53 Ticks and tick-borne pathogens provide a salient system for examining the effects of wildlife loss  
54 and climate changes on disease risk. Globally, ticks are considered to be the most important  
55 disease vectors for wildlife and domestic animals [10], and are second only to mosquitoes among  
56 vectors affecting humans [11]. Estimated economic costs of ticks and tick-borne disease are  
57 variable [12] and although no recent estimate has been made, one study attributed annual losses  
58 of US\$ 13.9 billion worldwide to tick-borne disease in cattle alone [13].

59 Globally, the pervasive decline in large-wildlife populations [14] is affecting a wide range of  
60 ecological functions and services, including disease control [15,16]. Ticks are also likely to be  
61 affected, considering their inextricable links to host population dynamics. While a substantial  
62 body of work demonstrates complex relationships among hosts, predators, and ticks (e.g., for the  
63 Lyme disease system in North America [17]), few studies have experimentally investigated how  
64 size-selective defaunation, which simulates the disproportionate vulnerability of larger animals  
65 to human disturbance [14], affects tick abundance and risk of tick-borne disease (but see [18]).  
66 Size-selective defaunation can directly affect tick abundance through the loss of hosts [19], and  
67 can also indirectly affect tick survival by altering vegetation structure [20–23] and the abundance  
68 and composition of small-vertebrate hosts [22,24]. Large-mammal loss often accompanies small-  
69 mammal abundance increases [22,24,25], leading to changes in host availability for different tick  
70 species. The relative importance of these sometimes opposing factors is poorly understood for  
71 most systems, and likely depends on vector life cycles and host associations.

72 Climate can also affect the prevalence and distribution of zoonotic pathogens, particularly those  
73 limited by climate-sensitive vectors [7,26–28]. This topic has become increasingly relevant in the  
74 context of global climate changes [7,9,29]. As tick survival can depend on factors such as rainfall  
75 and temperature [21,30,31], several models have predicted shifting tick ranges that result in net  
76 range expansions under climate change scenarios, although this varies among tick species [32].  
77 This experiment is one of few field studies that consider climatic effects on multiple tick species

78 simultaneously, and is situated in a region where climate changes are already pervasive and will  
79 be challenging to mitigate [33].

80 While the independent effects of climate change and biodiversity loss on zoonotic disease have  
81 received considerable recent attention, their potential interaction has not been well explored. For  
82 tick-borne diseases, prior studies have been largely correlative, yielding mixed results on the  
83 relative importance of various climate metrics, host abundance, and their interaction in  
84 determining tick abundance [34–37], emphasizing the need for more data describing a range of  
85 interacting forces on tick biology. The indirect effects of large herbivores on other small  
86 consumers, from insects to birds and small mammals, are highly sensitive to variation in climate  
87 and productivity [22,38,39], but it is not known whether these results can be generalized to  
88 disease risk in particular.

89 East African savannas are hotspots of tick and tick-borne pathogen diversity [40], and tick-borne  
90 pathogens such as *Rickettsia*, *Coxiella*, and *Anaplasma* are major regional economic and human  
91 health concerns [41–43]. For example, a recent study in Tanzania found that bacterial zoonoses  
92 caused 26% of acute fever cases; of these, 20% were Q Fever, caused by *Coxiella burnetii*, and  
93 30% were Rickettsiosis, caused by spotted fever group *Rickettsia* [44]. Accordingly, African  
94 savannas offer an ideal system for testing the effects of varying degrees of defaunation on tick  
95 abundance, as hosts are diverse and abundant, ranging over six orders-of-magnitude in size and  
96 occupying diverse functional roles [22,45]. However, large wildlife are experiencing widespread  
97 and precipitous declines in many parts of this region [46,47], underscoring the importance of  
98 predicting effects across ecological communities. Furthermore, climate change is also likely to  
99 affect tick-borne disease in East Africa, due in part to shifting rainfall patterns [31]. While large-  
100 scale predictions for future rainfall regimes are mixed [33], much of the region has been affected  
101 by persistent reductions in the critical ‘long rains’ since 1970 [48], and localized rainfall  
102 prediction models indicate that this trend is likely to continue [49].

103 We used a replicated series of experimental large-herbivore exclosures to quantify the effects of  
104 size-selective defaunation, climatic context, and their interaction on tick abundance and  
105 prevalence of tick-borne pathogens. In light of evidence that other consumer groups respond both  
106 numerically and behaviorally to an interaction between defaunation and primary productivity  
107 [38,39,50,51], we hypothesized that: (1) large-herbivore removal has strong effects on ticks and  
108 their associated pathogens; (2) tick species that utilize small mammal hosts will increase in  
109 abundance when large mammals are excluded (and small-mammal densities increase); and (3)  
110 the strength of these effects are contingent on climatic context and are strongest in more arid,  
111 low-productivity areas.

## 112 **Methods**

### 113 *Survey Site and Exclosures*

114 Research was conducted in the Ungulate Herbivory Under Rainfall Uncertainty (UHURU)  
115 experimental plots [22,52,53], established in 2008 at Mpala Research Centre (MRC) in Laikipia  
116 County, Kenya (0°17' N, 37°52' E, 1600m elevation). MRC supports robust populations of  
117 wildlife including elephants (*Loxodonta africana*), giraffe (*Giraffa camelopardalis*), zebra  
118 (*Equus grevyi* and *Equus quagga*), impala (*Aepyceros melampus*), and dik-dik (*Madoqua kirkii*),  
119 among others. The UHURU plots consist of four 1-ha exclosure treatments replicated three times  
120 at each of three 'levels' of a rainfall and productivity gradient created by the rain shadow of Mt.  
121 Kenya (i.e., 9 total replicates of each treatment, 36 total plots; Table S1). The four treatments  
122 simulate different scenarios of size-selective species losses using different combinations of  
123 fencing. The treatments are as follows: (1) total exclusion of all ungulate herbivores ('Total  
124 exclosure'); (2) exclusion of all herbivores >15kg ('Meso exclosure'); (3) exclusion of only  
125 mega-herbivores (i.e., giraffe and elephant; 'Mega exclosure'), and 4) unfenced open plots  
126 ('Control') [22]. Mean annual precipitation increases ~45% from the arid northern sites (440mm  
127 year<sup>-1</sup>), to the mesic southern sites (640mm year<sup>-1</sup>), with central sites intermediate (580mm year<sup>-1</sup>).  
128 Seasonal rains typically fall from March – May ('long rains') and October – December ('short  
129 rains') [54]. As in other semi-arid savannas, primary productivity is tightly linked to  
130 precipitation across this gradient [22]. Although the Normalized Difference Vegetation Index  
131 (NDVI) has been used previously in studies of tick abundance [21], we used mean annual rainfall  
132 as the primary climatic variable in our analyses, both because NDVI increases in exclosure  
133 treatments due to decreased herbivory and trampling by large mammals [22] (and thus would not  
134 isolate climatic factors), and because climatic factors tend to outperform NDVI in predicting  
135 African tick distributions [31]. We also present a complementary analysis using a categorical  
136 'climatic level' variable in lieu of the continuous precipitation variable; results are qualitatively  
137 similar (Tables S2, S3).

### 138 *Ticks*

139 The density of infected vectors is a common metric of vector-borne zoonotic disease risk  
140 [15,55,56] and is directly related to both vector density and pathogen infection rate. Thus,  
141 changes in tick density, infection rate, or a combination of the two can affect disease risk. To  
142 measure disease risk, we used tick drags and pathogen screening to quantify the density and  
143 infection rate of ticks.

### 144 *Tick Drags*

145 Ticks were collected in Total exclosure and Control plots each month for 13 months between  
146 October 2013 and November 2014. For each survey, a standard white canvas cloth was dragged  
147 throughout all passable portions of each plot, but areas of dense thicket areas were not sampled.  
148 Because exclosure plots often featured thick, thorny vegetation that precluded drags over fixed  
149 linear distances, we conducted drags for a 1-hour period, with ticks collected every 5 minutes.  
150 We also surveyed the Mega and Meso exclosure plots for five months in 2014 (Jan, July, Aug,

151 Sept, Nov). To ensure that drags accurately estimated the tick species composition of each plot,  
152 the drags were complemented with CO<sub>2</sub> traps [57] for two months.

153 Ticks were subsequently identified to species using microscopy and descriptions from [58]. We  
154 focused all analyses on three congeneric tick species—*Rhipicephalus pravus*, *R. praetextatus*,  
155 and *R. pulchellus*—that dominated the tick community. These tick species vary considerably in  
156 typical host preferences for each of their three distinct life stages (Figure S1). In general,  
157 immature stages of *R. pravus* and *R. praetextatus* feed upon small mammals (particularly  
158 rodents), which roughly double in abundance within total exclosures [22,53], whereas all stages  
159 of *R. pulchellus* feed on larger mammals [58,59]. Thus, the UHURU exclosure design alters the  
160 dominant host availability for each of these tick species (Figure S1; [22,53,58,59]).

### 161 *Pathogen Screening*

162 We extracted DNA and prepared double-indexed libraries for 136 ticks following [60]. Tick  
163 sample size was calculated to detect a 10% variation in pathogen prevalence across treatments  
164 while sampling across multiple species, treatments, and levels. Ticks with insufficient read data  
165 were excluded. Libraries were captured in pools of eight individuals (12.5ng each library per  
166 capture; 100ng total library per pool) using the Ectobaits protocol [60]. Double-indexed libraries  
167 were then amplified post capture with Illumina adapters by 18 cycles of PCR. Adapter multimers  
168 were removed prior to sequencing using QIAEX II Gel Extraction Kits (Qiagen). Captured  
169 products were sequenced on a MiSeq (Illumina, USA) using paired-end 150 bp reads. MiSeq  
170 library sequences underwent quality control as described in [60], except that minimum average  
171 base quality score was 25. We differentiated between *Coxiella burnetii* and *Coxiella*-like  
172 endosymbionts, as these groups are genetically similar, but endosymbionts are non-pathogenic  
173 and often have high infection rates [61]. We reanalyzed five libraries (KenT11b-KenT15b)  
174 included in [60]. For a subset of ticks (n=20), we confirmed *Rickettsia*, *Coxiella*, *Ehrlichia*, and  
175 *Anaplasma* infection and tick species using PCR assays following [60]. Positive PCR products  
176 were sequenced with an ABI 3130xl (Thermo Fisher Scientific, USA).

### 177 *Statistical Analyses*

178 We analyzed the tick drag data with generalized linear mixed models (GLMM), using counts of  
179 adult ticks per plot as our response variable [62]. Fixed effects included treatment (Total  
180 exclosure and Control for all months; all treatments for a subset of months), mean annual  
181 precipitation, and the treatment × rainfall interaction; random effects included replicate plot  
182 identity (3 plots within each of 3 rainfall levels; n=9) and time period (month; n = 12 for Total  
183 exclosure vs. Control, n=5 for all treatments). We ran two separate sets of GLMMs, one for Total  
184 exclosure and Control plots across all months, and another for all plots for the subset of five  
185 months. Candidate-model sets included all possible combinations of the two main effects and  
186 their interaction (the “full model”), along with a null model; all models included the random  
187 effects (Table 1, Table S4). We analyzed the combined total of all tick species and each species  
188 separately. As data were overdispersed and zero-inflated for individual tick species, we used

189 zero-inflated negative-binomial distributions with log link functions in our GLMMs. For the two  
190 datasets that combined the three tick species, we used negative-binomial distributions with log  
191 link functions. All models were constructed using the glmmADMB package in R [63,64].

192 All model combinations for each tick species and the combined total of ticks were ranked using  
193 the second-order Akaike's information criterion (AICc) [62] using the MuMIn package [65]. We  
194 investigated all models (reported in S5 and S6), and present the 95% confidence interval set with  
195 individual parameter estimates and Akaike weights ( $w_i$ ) in Tables 1 and 2.

196 *Coxiella burnetii* and *Rickettsia* spp. were the only pathogens sufficiently prevalent to permit  
197 robust statistical analysis. We analyzed the likelihood of infection using binomial GLMMs with  
198 logit link functions, with infection status of each tick (infected/uninfected) as the response.  
199 Experimental treatment, tick species, rainfall, and treatment  $\times$  rainfall were fixed effects and plot  
200 replicate was a random effect.

201 All analyses were performed in R version 3.3.0 [66]. Descriptive statistics are reported as mean  
202 number of ticks per ha  $\pm$  1 standard error.

## 203 **Results**

204 In total, we captured 5677 ticks across all plots, including 4180 via tick drags and 1497 via traps.  
205 Of these, >95% were adults of just three species: *R. pravus* (43%), *R. praetextatus* (36%), and *R.*  
206 *pulchellus* (17%). Adults were substantially more abundant than other life stages in both drag  
207 and trap collections, despite efforts to avoid under-sampling juvenile ticks. Fewer than 3% of the  
208 ticks captured were nymphs, and no larvae were collected. Tick traps did not capture additional  
209 tick species; therefore, we used only drag data for all subsequent analyses (S1, Table S4, and  
210 Figures S2 and S3) and focused all analyses on adults of the three dominant species.

### 211 *Total abundance of the three dominant tick species*

212 Total tick abundance varied seasonally over the 13-month sampling period and the scale and  
213 timing of fluctuations differed among tick species (Figure 1A). However, on average, total tick  
214 abundance doubled in Total exclosures ( $18.3 \pm 1.9$ ) relative to Control plots ( $9.9 \pm 1.0$ ) (Figure 1A,  
215 B, Table 1). Low-rainfall plots had 225% more ticks on average ( $17.8 \pm 2.3$ ) than did mesic plots  
216 ( $7.9 \pm 1.0$ ). Total tick abundance was best explained by the GLMM that included exclosure  
217 treatment, precipitation, and their interaction (Tables 1, S5) ( $w_i = 0.75$ ). The interaction ( $z = -$   
218  $2.3$ ,  $P=0.02$ ; Table 1) reflected the increasing effect of wildlife exclusion on tick abundance as  
219 aridity increased (Figure 1C; Tables 1, S5). We found some support ( $w_i = 0.16$ ) for a model with  
220 no interaction and a marginally-negative relationship between rainfall and tick abundance ( $z = -$   
221  $1.96$ ,  $P=0.05$ ). Net results were similar in the analysis that considered all four wildlife-exclusion  
222 treatments for a subset of months: total tick abundance increased from 170% (only  
223 megaherbivores excluded) to 360% (all large wildlife excluded) (Figure 1D). The full model was  
224 again the best fit ( $w_i = 0.99$ ), with significant interactions between rainfall and the Total and  
225 Meso exclosure treatments ( $z = -3.61$ ,  $P=0.001$ , Total;  $z = -3.38$ ,  $P=0.001$ , Meso; Tables 2, S6).

## 226 *Species-specific responses*

227 Although *R. pravius* and *R. praetextatus*, two tick species that often parasitize smaller mammals,  
228 increased with large mammal loss, only *R. pravius* abundance showed clear evidence of an  
229 interaction between enclosure and aridity. For the full 13 months of data, the best model for *R.*  
230 *pravius* included treatment, rainfall, and their interaction ( $w_i = 0.99$ ), whereas the best model for  
231 *R. praetextatus* included only treatment ( $w_i = 0.47$ ) and a second model ( $w_i = 0.37$ ) included the  
232 non-significant effect of rainfall (Table 1). Both tick species increased in Total enclosures  
233 relative to Controls ( $z = 8.40$ ,  $P < 0.001$ , *R. pravius*;  $z = 3.74$ ,  $P < 0.001$ , *R. praetextatus*), and this  
234 effect was stronger in drier sites for *R. pravius* only ( $z = -3.37$ ,  $P < 0.001$ ). By contrast, rainfall had  
235 no detectable effect on tick abundance in Control plots ( $z = -1.02$ ,  $P = 0.31$ ). For the subset of data  
236 collected in all four wildlife exclusion treatments, the full model was the best fit for both tick  
237 species ( $w_i = 0.93$ ,  $w_i = 0.78$ , *R. pravius* and *R. praetextatus* respectively). Both tick species  
238 increased in all enclosure treatments relative to Controls, and both increased significantly in  
239 Total enclosures ( $z = 7.22$ ,  $P < 0.001$ ;  $z = 4.07$ ,  $P < 0.001$ ) (Table 2). This effect was more  
240 pronounced in drier sites for both species, although this was only significant for *R. praetextatus*  
241 in Meso enclosures ( $z = -2.26$ ,  $P = 0.02$ ) and *R. pravius* in Total enclosures ( $z = -3.26$ ,  $P < 0.001$ )  
242 (Table 2.). A second model for *R. praetextatus* that included only treatment ( $w_i = 0.13$ ) received  
243 considerably less support.

244 For *R. pulchellus*, which often parasitize larger-bodied mammals, the best model for all months  
245 included only enclosure treatment ( $w_i = 0.48$ ), and a second model ( $w_i = 0.39$ ) included the non-  
246 significant effect of rainfall; but here Total wildlife exclusion caused a 43% decrease in  
247 abundance relative to Controls ( $z = -1.95$ ,  $P = 0.05$ ; Table 1, Figure 1B). For the subset of data  
248 including all four treatments, the best model ( $w_i = 0.46$ ) again included only enclosure treatment,  
249 while a second model ( $w_i = 0.37$ ) included the non-significant effect of rainfall. However, this  
250 secondary analysis revealed that partial wildlife exclusion caused increases in tick abundance  
251 relative to controls ( $z = 4.72$ ,  $P < 0.001$ , Meso;  $z = 2.44$ ,  $P = 0.02$ , Mega; Tables 2, S6, Figure 1D),  
252 but total exclusion had no significant effect ( $z = -0.57$ ,  $P = 0.57$ ).

## 253 *Pathogens*

254 Prevalence of *C. burnetii* isolates was 43% ( $n = 58$  of 136 ticks screened), and prevalence of  
255 *Rickettsia spp.* was 5% ( $n = 7$  of 136 ticks; 4 of these were from the spotted fever group). We  
256 detected *Ehrlichia* in 1 adult tick and *Anaplasma* in 1 nymph (nymphs were not analyzed due to  
257 small sample size). We found a high prevalence of non-pathogenic *Coxiella*-like endosymbionts  
258 (57%; 46% of these were also present in ticks with confirmed *C. burnetii* isolates). Therefore,  
259 our analyses excluded ticks for which only a *Coxiella*-like endosymbiont was detected, but  
260 included ticks with both *C. burnetii* and *Coxiella*-like endosymbionts.

261 For the GLMM for *C. burnetii*, no combination of our predictors outperformed the null model  
262 (Table 3). For *Rickettsia spp.*, the best model included tick species and rainfall; however, neither  
263 estimate was significant (although rainfall marginally increased infection probability; Table 3).



264 In sum, there were no pronounced effects of treatment, tick species, or rainfall on pathogen  
265 prevalence (Figure 2, Tables 3, S7, Figure S4).

## 266 **Discussion**

267 Our results support our hypothesis that defaunation and climate can interact to markedly affect  
268 the abundance of ticks and thus the risk of tick-borne disease exposure (although not necessarily  
269 the prevalence of these pathogens). Total exclusion of all large wildlife increased total tick  
270 abundance by 130% (mesic sites)-225% (arid sites), showing a significant interaction with  
271 aridity. Tick abundance increased from 170% (only mega-herbivores excluded) to 360% (all  
272 large wildlife excluded) during the five-month period in which all enclosure plots were surveyed.  
273 We found no significant variation in pathogen prevalence across plots or tick species, suggesting  
274 that the risk of tick-borne pathogen exposure reflects observed tick abundance patterns.

275 However, this overall pattern masks strong differences in the magnitude and direction of effects  
276 of wildlife exclusion across tick species and over time. Tick species-specific responses show  
277 some overlap with expectations based on tick host associations. Patterns in total tick abundance  
278 were driven by two dominant tick species, *R. pravius* and *R. praetextatus*, whose immature stages  
279 frequently feed upon small hosts, which also increase strongly following wildlife exclusion  
280 [22,50,51]. Although we do not expect changes in adult tick abundance to directly correlate with  
281 fluctuations in rodent abundance in these plots over time, a comparison of long-term rodent  
282 abundance and tick abundance within each plot produces positive correlations for *R. pravius* and  
283 *R. praetextatus* ( $z = 6.59$ ,  $P < 0.001$  and  $z = 3.17$ ,  $P < 0.01$ , respectively; Table S8). In contrast, the  
284 third common tick species, *R. pulchellus*, whose adult stages primarily parasitize vertebrates  
285 larger than 15kg [58], and whose immature stages are not found on rodents [59], decreased with  
286 the total absence of large wildlife for the 13-month dataset. However, for the five months for  
287 which all four enclosure treatments were surveyed, abundance of this tick species in total  
288 enclosures was no different from that in controls, but we observed marked increases in  
289 abundance within partial wildlife enclosures (see Figure S1 for tick/host associations in  
290 enclosure plots). This discrepancy highlights temporal variation in enclosure effects: strong  
291 changes occur during months of peak tick abundance, which were not captured by the five-  
292 month dataset.

293 Other factors beyond the release of intermediate hosts may have also influenced the marked  
294 differences in adult tick abundance among experimental plots. Increases in small carnivores  
295 (potential hosts for all three tick species) in response to elevated rodent density in enclosure plots  
296 may increase total tick abundance [18]. Likewise, increases in understory vegetation cover  
297 following large wildlife loss may increase tick survivorship (via lowered risk of desiccation)  
298 [22]. The relative importance of these factors may vary among tick species depending on their  
299 life histories. The complex pathways by which wildlife loss may affect the abundance of  
300 different tick species likely explains why the few previous studies on the effects of large wildlife  
301 enclosure on tick abundance have produced mixed results [18,67].

302 Total tick abundance was greater in drier areas, although this pattern was largely driven by the  
303 most common tick species, *R. pravus*. *Rhipicephalus praetextatus* and *R. pulchellus* only  
304 increased modestly in these areas, and annual rainfall was not a major explanatory factor in  
305 models of their abundance. This is consistent with previous observations of climate preferences  
306 for these species, as *R. pravus* may particularly favor areas with extended dry seasons [61].  
307 Notably, tick community composition varied considerably over seasons, and the most significant  
308 responses to exclosures occurred at months of peak abundance (Figure 1A). These months of  
309 peak abundance drove overall patterns for each species, and are likely to be a result of strong  
310 differences in tick phenology and responsiveness to rainfall.

311 *Rhipicephalus pravus* also drove an interaction between wildlife-exclosure treatment and aridity  
312 on tick abundance, despite variation among tick species. This interaction and its variation are  
313 consistent with prior studies of the effects of defaunation on consumer communities, including a  
314 recent meta-analysis that found these effects are often context-dependent and mediated by site  
315 productivity [39,50,68]. In this region, rodent-borne pathogens have shown a similar response:  
316 anthropogenic disturbance tends to cause stronger increases in rodent-borne disease in drier  
317 climates with lower productivity [69]. However, consistent with our findings here, responses are  
318 variable across specific hosts and pathogens [69].

319 Both pathogens analyzed in this study are globally important. *C. burnetii*, the causative agent of  
320 Q Fever, is considered to be an emerging zoonotic disease [70], while rickettsial pathogens are  
321 responsible for a variety of spotted fevers—including African tick-bite fever (caused by  
322 *Rickettsia africae*) in our study location [42]. We observed no significant differences in  
323 prevalence of either *C. burnetii* or *Rickettsia* spp. due to wildlife exclosure treatment, rainfall, or  
324 tick species. Larger sample sizes and screening over many seasons might reveal finer-scale  
325 dynamics; however, on a coarse level, this result suggests that tick-borne disease risk is likely to  
326 be well-approximated by estimates of total tick abundance (Figure 2). *Coxiella burnetii*  
327 prevalence was surprisingly high. Although we excluded ticks for which only an endosymbiont  
328 was detected, 67% of the ticks infected with *C. burnetii* were also positive for the *Coxiella*-like  
329 endosymbiont. Endosymbionts may benefit some ticks [61], and recent work suggests that *C.*  
330 *burnetii* recently emerged from this group [71]. Thus, the genetic similarity between *C. burnetii*  
331 and *Coxiella*-like endosymbionts may have yielded some false positives given that the full  
332 *Coxiella* phylogeny is incomplete. However, we do not expect this to bias our results, given that  
333 the likelihood of false positives is consistent across all predictors.

334 Our study demonstrates the significant potential for size-selective defaunation to alter the risk of  
335 tick-borne disease. Substantial variation in tick abundance and species composition over time  
336 reflect the inherent complexity of a system that depends on host, environmental, and vector  
337 variables, but total effects suggest long-term patterns, especially when ticks peak in abundance.  
338 On average, when all large wildlife were excluded, the total number of ticks nearly doubled; and,  
339 when only Mega wildlife and Meso wildlife were excluded (perhaps a more realistic short-term  
340 defaunation scenario for much of the world), ticks of all three major species increased,

341 suggesting that large-wildlife loss can contribute to an increased tick-borne disease risk that may  
342 be mitigated by conservation in many contexts. Furthermore, the costs of wildlife loss on tick-  
343 borne disease in this region may be intensified in drier, less productive areas that are likely to  
344 worsen with a changing climate [48], demonstrating interacting effects of wildlife loss and  
345 climate change on tick-borne disease risk. On a more global scale, our study highlights the  
346 challenge of predicting the effects of either biodiversity loss or climate change in isolation of  
347 other stressors on vector ecologies and infectious disease dynamics.

348 Author contributions:

349 GCT conducted fieldwork, performed analyses, and wrote the first version of the manuscript;  
350 BFA and TH identified ticks and contributed to the final report; TA conducted fieldwork, tick  
351 identification, sample logistics, and data-entry; LH conducted all biomolecular lab work; RMP  
352 and TMP designed and provided access to experimental plots, contributed data, and assisted with  
353 data interpretation and writing the final report; LN assisted in obtaining Kenyan research permits  
354 and provided report feedback; MGC conducted bioinformatic analyses and provided report  
355 feedback; RF coordinated and supervised molecular work; JNM conducted fieldwork; HSY  
356 conceived the project and analyses, and assisted with data interpretation and writing the  
357 manuscript.

358 Competing interests:

359 We declare no competing interests.

360 Data availability:

361 Datasets and R code for all analyses are available at: <https://github.com/gtitcomb/Wildlife-loss-climate-ticks>.

363 Read data from this project are available in the BioProject Archive, accession PRJNA362357.  
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Figure 1: (A) Tick abundance varied over time, across rainfall levels, among species, and between treatments for the full 13-month dataset. (B) While total tick abundance increased in Total exclusions, the magnitude and direction of this effect varied by tick species for the 13-month dataset. (C) For all tick species summed together, exclusion interacted with annual rainfall, with stronger effects of exclusion in drier environments. (D) When all exclusions were surveyed for the 5-month subset of data, tick species responded differently to varied wildlife loss levels. Asterisks indicate significant ( $P < 0.05$ ) differences from Control plots (green); dots indicate non-significant trends ( $P < 0.1$ ).

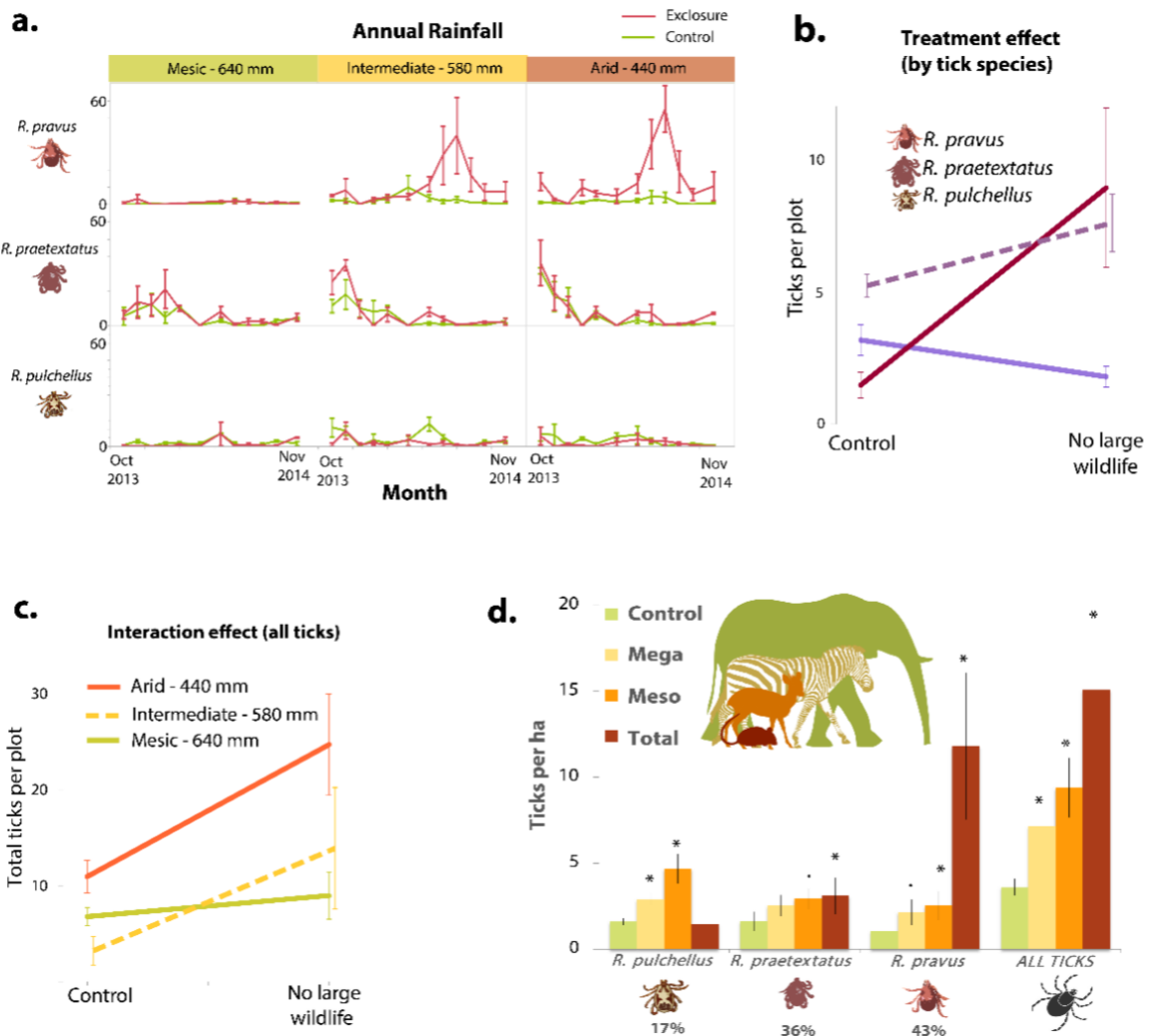
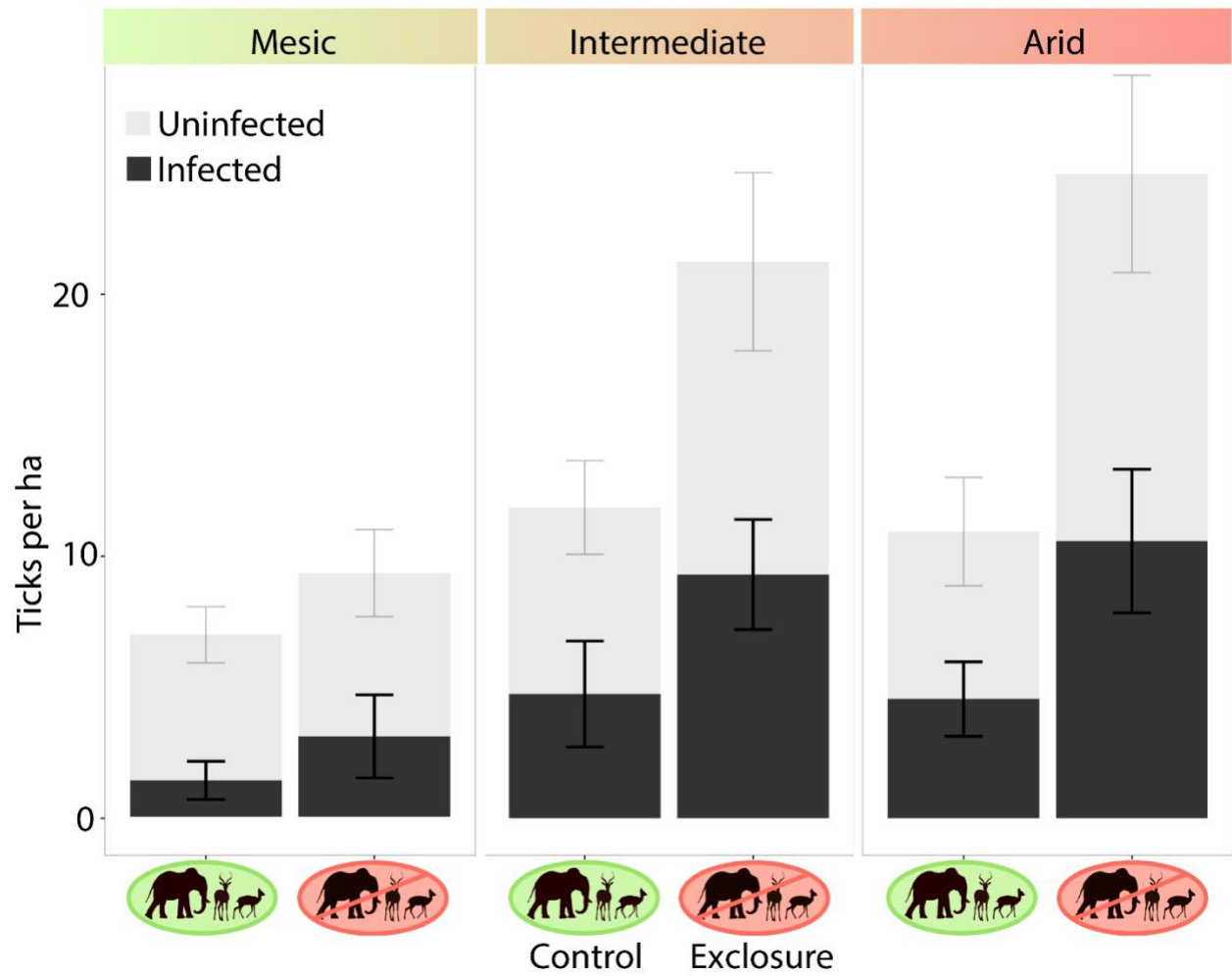


Figure 2: The estimated number of infected and uninfected ticks increased in plots where large wildlife had been removed (Exclosure), and this was further increased in arid sites, regardless of tick-borne pathogen infection.



Model	Intercept		Exclosure		Rainfall		Exclosure × Rainfall	
<b>All Ticks</b> $w_i = 0.75$	2.117 ± 0.200	10.61 <0.001	<b>0.587 ± 0.130</b>	4.53 <0.001	-0.092 ± 0.142	-0.65 0.52	<b>-0.295 ± 0.128</b>	-2.30 0.02
<b>R. pravus</b> $w_i = 0.99$	-0.173 ± 0.388	-0.44 0.66	<b>1.452 ± 0.173</b>	8.40 <0.001	-0.304 ± 0.297	-1.02 0.31	<b>-0.596 ± 0.177</b>	-3.37 <0.001
<b>R. praetextatus</b> $w_i = 0.47$	1.157 ± 0.441	2.63 0.009	<b>0.431 ± 0.115</b>	3.74 <0.001				
<b>R. pulchellus</b> $w_i = 0.47$	0.896 ± 0.312	2.87 0.004	<b>-0.441 ± 0.227</b>	-1.95 0.05				

Legend:  

Estimate ± SE	z-score
	P-value

Table 1: Effects of exclosure treatment, rainfall, and their interaction for all months (Control and Total exclosure plots only) from four GLMMs. Control plots are designated as the reference, and rainfall (mm) is scaled by standard error (84 mm) and centered at the mean (533 mm) for ease of interpretation. Significant relationships ( $P < 0.05$ ) are bolded. Positive relationships are shaded in blue; negative relationships are shaded yellow. All estimates are shown with standard errors, z-score (upper right), and P-value (lower right). Full model sets and parameters are shown in Table S5.

Model	Intercept		Exclosure		Rainfall		Exclosure × Rainfall	
<b>All Ticks</b> $w_i = 0.99$	$1.424 \pm 0.232$	6.14 <0.001	TOTAL	6.92	$0.219 \pm 0.162$	1.35 0.176	TOTAL	-3.61
			<b><math>1.161 \pm 0.168</math></b>	<0.001			<b><math>-0.62 \pm 0.2</math></b>	<0.001
			MESO	4.11			MESO	-3.38
			<b><math>0.698 \pm 0.170</math></b>	<0.001			<b><math>-0.585 \pm 0.173</math></b>	<0.001
			MEGA	3.01			MEGA	-0.87
			<b><math>0.511 \pm 0.170</math></b>	0.003			$-0.150 \pm 0.173$	0.39
<b>R. pravus</b> $w_i = 0.93$	$-0.511 \pm 0.520$	-0.98 0.33	TOTAL	7.22	$-0.228 \pm 0.365$	-0.62 0.53	TOTAL	-3.26
			<b><math>1.904 \pm 0.264</math></b>	<0.001			<b><math>-0.874 \pm 0.268</math></b>	0.001
			MESO	2.23			MESO	-1.49
			<b><math>0.627 \pm 0.282</math></b>	0.03			$-0.428 \pm 0.288$	0.14
			MEGA	1.89			MEGA	-0.91
			$0.537 \pm 0.283$	0.06			$-0.267 \pm 0.295$	0.37
<b>R. praetextatus</b> $w_i = 0.78$	$0.159 \pm 0.516$	0.31 0.76	TOTAL	4.07	$0.499 \pm 0.268$	1.86 0.062	TOTAL	-1.26
			<b><math>1.050 \pm 0.258</math></b>	<0.001			$-0.350 \pm 0.278$	0.208
			MESO	1.70			MESO	-2.26
			$0.448 \pm 0.263$	0.09			<b><math>-0.628 \pm 0.278</math></b>	0.024
			MEGA	1.10			MEGA	-0.10
			$0.292 \pm 0.266$	0.27			$-0.030 \pm 0.286$	0.92
<b>R. pulchellus</b> $w_i = 0.46$	$0.463 \pm 0.200$	2.32 0.021	TOTAL	-0.57				
			$-0.155 \pm 0.273$	0.57				
			MESO	4.72				
			<b><math>1.184 \pm 0.251</math></b>	<0.001				
			MEGA	2.44				
			<b><math>0.630 \pm 0.258</math></b>	0.015				

Legend:

<b>Estimate ± SE</b>	z-score
	P-value

Table 2: Effects of all exclosure treatments on tick abundance (for a subset of months) from four GLMMs. Exclosure compares Control plots (all wildlife allowed), the reference, to plots that selectively exclude mega herbivores (MEGA), mega and meso herbivores (MESO), and all herbivores greater than 5kg (TOTAL). Rainfall (mm) is scaled by standard error (84 mm) and centered at the mean (533 mm) for ease of interpretation. Significant relationships ( $P < 0.05$ ) are bolded, marginally significant relationships ( $P < 0.1$ ) are bordered by a broken line, positive relationships are shaded blue, and negative relationships are yellow. All estimates are shown with standard errors, z-score (upper right), and P-value (lower right). Full model sets and parameters are shown in Table S6.

	Intercept		Species	Exposure	Rain	Exposure x Rain
<b><i>C. burnetii</i></b>						
<b>Model 1</b>						
$w_i = 0.26$	$-0.42 \pm 0.2$	-2.38				
		0.02				
<b><i>Rickettsia</i> sp.</b>						
<b>Model 1</b>						
$w_i = 0.35$			RHPU	-0.02		
				$-14.8 \pm 799$	0.99	
	$-2.44 \pm 0.7$	-3.48				$1.16 \pm 0.6$ 1.87
		<0.001	RHPV	-1.53		0.06
				$-1.25 \pm 0.8$	0.13	
<b>Model 2</b>						
$w_i = 0.18$						
	$-3.37 \pm 0.6$	-5.74				$1.09 \pm 0.6$ 1.87
$\Delta AIC_c = 1.36$		<0.001				0.06

Table 3: Results of GLMMs for *Coxiella burnetii* and *Rickettsia* sp. ‘Species’ compares the probability of tick infection with each pathogen for each tick species (*R. pravus* – RHPV, and *R. pulchellus* – RHPU, as compared to *R. praetextatus* – RHPR). Marginally significant relationships ( $P < 0.1$ ) are bordered by a broken line. All estimates are shown with standard errors, z-score (upper right), and P-value (lower right).